



# Probability of Occurrence and Displacement Regression of Distributed Surface Rupturing for Reverse Earthquakes

*Fiia Nurminen*<sup>1\*</sup>, *Paolo Boncio*<sup>1,2</sup>, *Francesco Visini*<sup>3</sup>, *Bruno Pace*<sup>1,2</sup>, *Alessandro Valentini*<sup>1</sup>, *Stéphane Baize*<sup>4</sup> and *Oona Scotti*<sup>4</sup>

<sup>1</sup>DiSPUTer, University “G. d’Annunzio” of Chieti-Pescara, Chieti, Italy, <sup>2</sup>CRUST—Centro interUniversitario per l’analisi SismoTettonica tridimensionale con applicazioni territoriali, Chieti, Italy, <sup>3</sup>INGV—Istituto Nazionale di Geofisica e Vulcanologia, Sezione di Pisa, Pisa, Italy, <sup>4</sup>IRSN—Institut de Radioprotection et de Sûreté Nucléaire, Fontenay-aux-Roses, France

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### \*Correspondence:

Fiia Nurminen  
fiia.nurminen@unich.it

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Probabilistic fault displacement hazard analysis provides a systematic approach to estimate the likelihood of occurrence and expected amount of surface displacement during an earthquake on-fault (principal fault rupturing) and off-fault (distributed rupturing). The methodology is based on four key parameters describing the probability of occurrence and the spatial distribution of the displacement both on and off-fault. In this work we concentrate on off-fault rupturing, and develop an original probability model for the occurrence of distributed ruptures and for the expected displacement distribution based on the compilation and reappraisal of surface ruptures from 15 historical crustal earthquakes of reverse kinematics, with magnitudes ranging from  $M_w$  4.9 to 7.9. We introduce a new ranking scheme to distinguish principal faults (rank 1) from simple distributed ruptures (rank 2), primary distributed ruptures (rank 1.5), bending-moment (rank 21) and flexural-slip (rank 22) and triggered faulting (rank 3). We then used the rank 2 distributed ruptures with distances from the principal fault ranging from 5 to 1,500 m. To minimize bias due to the incomplete nature of the database, we propose a “slicing” approach as an alternative to the “gridding” approach. The parameters obtained from slicing are then combined with Monte Carlo simulations to model the dependence of the probability of occurrence and exceedance with the dimensions and position of the site of interest with respect to the principal fault, both along and across strike. We applied the probability model to a case-study in Finland to illustrate the applicability of the method given the limited extend of the available dataset. We finally suggest that probabilistic fault displacement hazard model will benefit by evaluating spatial distribution of distributed rupture in the light of spatial completeness of the input data, structural complexity and physics observables of the causative fault.

**Keywords:** probabilistic fault displacement hazard analysis, surface faulting, distributed rupturing, displacement regression, reverse earthquakes

## INTRODUCTION

Probabilistic fault displacement hazard analysis (PFDHA) is a method developed for characterizing the expected amount and distribution of co-seismic fault displacement at the surface. In a PFDHA, the probability of exceeding a certain amount of displacement at a site is represented as a function of the earthquake magnitude frequency distribution, the ensuing probability of surface rupturing, and distribution of distances from the fault trace to the site. The key element in surface rupturing analysis is the differentiation between the principal fault rupturing (PF), and distributed rupturing (DR). PF is the fault plane along which the energy releases during the seismic event; DR (also known as secondary rupturing) refers to all the other surface ruptures, often less continuous, occurring away from PF. The PFDHA methodology was developed for normal faulting environments by the working group of Youngs et al. (2003) and developed further for strike-slip faults by Petersen et al. (2011). For the case of reverse faults, Moss and Ross (2011) worked specifically on PF data whereas Takao et al. (2013) worked on both strike-slip and reverse faults but using data only from Japanese earthquakes. In depth analysis of the DR data for reverse faulting earthquakes has not received much attention from the scientific community. The present work is a first step toward filling this gap.

In the following, this article describes the construction of the database of surface ruptures for reverse faulting earthquakes, the methodology used for obtaining the empirical parameters, and the mathematical approach used to model the occurrence and displacement distribution of DR based on the database. In the final part, the new regression parameters are applied to model the distributed surface rupturing hazard along the Suasselkä post-glacial fault (PGF) situated in Lapland, Northern Finland. It is a several tens of kilometres long reverse fault along which several geological and geophysical research projects are ongoing. This fault was chosen as an example case because the Suasselkä PGF may be considered still active today (Ojala et al., 2019a) and thus a potential source of a surface rupturing earthquake.

## CONSTRUCTING THE DATABASE AND RANKING THE OBSERVATIONS

The work initiated by developing a database of surface faulting data of well documented reverse faulting earthquakes. The displacement parameters and the spatial distribution of surface faulting were analyzed from georeferenced maps and published datasets. Following the SURE database architecture (Baize et al., 2019), we built a database of surface rupturing displacement reported in publications for 15 historical medium-large reverse earthquakes of moment magnitude from  $M_w$  4.9 to  $M_w$  7.9, occurred between 1970 and 2019. The selected events were all studied with an acceptable accuracy on site, and there are good quality displacement data and mapped fault traces available. All the displacement measurements and fault orientation parameters reported by the first-hand authors were entered in the database as published. Although displacement data available is not uniform, all the published information was entered as it was published

introducing the least interpretation possible. The events, their key parameters used in this study, and the referenced documents used are listed in **Table 1**.

The ruptures in the database were distinguished between principal fault rupture (PF) and DR following the definition of Youngs et al. (2003) where the PF tend to have rather continuous fault traces hosting the major displacements, whereas DR are characterized often by shorter, discontinuous traces of smaller displacement in respect to the PF. However, in order to be able to clearly identify and possibly exclude data affected by structural complexities in the derivation of PFDHA empirical regressions, several sub-ranking categories for DR are also proposed beyond the classical PF and DRs categories (**Figure 1**). Distinguishing between these is seen as important due to the different mechanisms behind them (Boncio et al., 2018). Consequently, for the generalized analysis, we are able to exclude the complex-type DR, that are likely related to the local structural setting, namely the occurrence of active folding related to the earthquake as illustrated in **Figure 1**. In the simplest case, rupturing along PF (rank 1) provokes spontaneous and discontinuous rupturing beyond this PF with a displacement significantly smaller which is referred to as “simple DR” (rank 2). By this we mean the rupturing occurred as a direct response to the surface rupturing along the PF in unpredictable locations. In favorable conditions, surface rupturing may take place also on pre-existing faults connected to the PF at depth along so called “primary DR” (rank 1.5), or along a disconnected, distant triggered fault (rank 3). Primary DRs may actually host significant displacement, and their surface traces are often relatively continuous. Triggered surface faulting takes place typically along pre-existing fault at a remarkable distance from the PF, but the surface fault is usually rather discontinuous. Furthermore, reverse faulting may lead to bending-moment (rank 21) and flexural-slip (rank 22) DR as defined by Philip and Megharoui (1983) and Yeats (1986); both of these being highly influenced by the specific geological structures favoring this type of faulting (especially active folding in relation to earthquake faults). The database contents by earthquake and fault ranking are listed in **Supplementary Table S2**.

The dataset is based on empirical field work, and the locations of the earthquakes and their surface deformation zones vary from very urbanized to hardly reachable mountain areas. During field survey after a surface-rupturing earthquake, geological investigations mainly focus on tracing and measuring PFs. Moreover, surface extension and spatial density of the DR are not always uniform along the ruptures of the corresponding PF and there is a large variability among the 15 earthquakes. We can attribute the variability in the database to a) the spatial completeness of data available, or b) physical reasons. For example, the spatial completeness can be affected by some areas not being accessible for field surveying, or the surface extension of the PF was too large to map all the DRs in detail. Also, the length of the DRs is a subject of mapping accuracy and interpretation. For example, in areas like San Fernando (urban area), many DRs are measured along the streets, thus rupture length  $\sim$  street width, and single traces are drawn discontinuous under buildings and similar obstacles. Physical elements that may have caused bias in the data coverage are due to DR formation

**TABLE 1** | Reverse earthquakes included in the database.

Earthquake	Date (yyyymmdd) = EventID	Moment magnitude ( $M_w$ )	Kin <sup>a</sup>	SRL <sup>b</sup> (km)	MD <sub>v</sub> <sup>c</sup> (m)	MD <sub>n</sub> <sup>d</sup> (m)	Types of DR <sup>e</sup>	References
Calingiri, Australia	19700310	5.0	RL-L	3.3	0.4	1.2	2	1, 2
San Fernando, CA, United States	19710209	6.6	R-LL	16	0.76	2.5	2, 1.5, 21, 3	1, 3, 4
El Asnam, Algeria	19801010	7.1	R	31	5.0	6.5	2, 21, 22, 3	1, 5–7
Coalinga (Nunez), CA, United States	19830611	5.4	R	3.3	0.50	1.0	2	1, 8
Marryat Creek, Australia	19860330	5.8	R-LL	13	0.9	1.1	2	1, 9–11
Tennant Creek, Australia	19880122 (3 events)	6.3	R	10.2	1.20	2.84	2	1, 12, 13
		6.4	R-LL	6.7	1.10	2.60	2, 3	
		6.6	R	16	1.77	2.50	2	
Spitak, Armenia	19881207	6.8	R-RL	25	1.6	2	2	1, 14, 15
Killar, India	19930929	6.2	R	5.5	0.6	1.2	2	16–18
Chi Chi, Taiwan	19990920	7.6	R-LL	72	9.8	11.65	2, 1.5, 3	19–32 <sup>f</sup>
Kashmir, Pakistan	20051008	7.6	R	70	3.40	7.05	2, 21	33–35
Wenchuan, China	20080512	7.9	R-RL	312	6.9	13.0	2, 1.5, 3	36–48
Pukatja, Australia	20120323	5.4	R	1.6	0.5	1.0	2	49
Nagano, Japan	20141122	6.2	R	9.34	0.8	1.60	2	50–53
Petermann, Australia	20160520	6.1	R-LL	20	0.9	1.9	2	54
Le Teil, France	20191111	4.9	R	5	0.23	0.33	2	55

<sup>a</sup>Kin. (kinematics): R, reverse; LL, left-lateral; RL, right-lateral.

<sup>b</sup>SRL, surface rupture length (principal faulting).

<sup>c</sup>MD<sub>v</sub>, maximum vertical displacement.

<sup>d</sup>MD<sub>n</sub>, maximum net displacement.

<sup>e</sup>Types of distributed rupturing: 2, simple; 1.5, primary distributed; 21, bending-moment; 22, flexural-slip; 3, sympathetic.

<sup>f</sup>For Chi Chi, we used the data from Lin (2000), Chen et al. (2000), Huang et al. (2008), Kelson et al. (2001), Huang et al. (2008), Angelier et al. (2003), Kelson et al. (2003), Ota et al. (2007), and Lee et al. (2003); Table 1, the points shown in maps Figures 4, 5. The points considered cover only the central-northern part of the rupture, but we do not expect a major bias in results of this work as, according to the reference maps, most of the distributed rupturing took place in northern section.

References: 1, Wells and Coppersmith, 1994; 2, Gordon and Lewis, 1980; 3, U.S. Geological Survey Staff, 1971; 4, Kamb et al., 1971; 5, Yelding et al., 1981; 6, Philip and Meghraoui, 1983; 7, Meghraoui et al., 1988; 8, Rymer et al., 1983; 9, Fredrich et al., 1988; 10, Bowman and Barlow, 1991; 11, Machette et al., 1993; 12, McCaffrey, 1989; 13, Crone et al., 1992; 14, Haessler et al., 1992; 15, Philip et al., 1992; 16, Lettis et al., 1997; 17, Seeber et al., 1996; 18, Rajendran et al., 1996; 19, Wesnousky, 2008; 20, Kelson et al., 2001; 21, Kelson et al., 2003; 22, Angelier et al., 2003; 23, Bilham and Yu, 2000; 24, Chen et al., 2000; 25, Huang et al., 2008; 26, Huang et al., 2000; 27, Lee et al., 2001; 28, Lee et al., 2003; 29, Lee et al., 2010; 30, Lin, 2000; 31, Ota et al., 2007; 32, Central Geological Survey, MOEA, 2017; 33, Avouac et al., 2006; 34, Kaneda et al., 2008; 35, Kumahara and Nakata, 2007; 36, Xu et al., 2009; 37, Liu-Zeng et al., 2009; 38, Liu-Zeng et al., 2012; 39, Yu et al., 2009; 40, Yu et al., 2010; 41, Zhou et al., 2010; 42, Zhang et al., 2013; 43, Chen et al., 2008; 44, Liu-Zeng et al., 2010; 45, Wang et al., 2010; 46, Xu et al., 2008; 47, Zhang et al., 2012; 48, Zhang et al., 2010; 49, King et al., 2019; 50, Okada et al., 2015; 51, Ishimura et al., 2015; 52, Lin et al., 2015; 53, Ishimura et al., 2019; 54, King et al., 2018; 55, Ritz et al., 2020.

being favored in correspondence to geometrical complexities connected to the PF or reflect the co-seismic patches at depth. For addressing the issue of likely incompleteness of the DR mapping, we developed an approach that generalizes the probability of DR occurrence along the strike of PF by making no assumption of the completeness of the database; we call this approach “slicing” and it is explained more in detail in the following chapter. Later, we balance this probability with the ratio of the total length of DR with respect to the PF length for each earthquake.

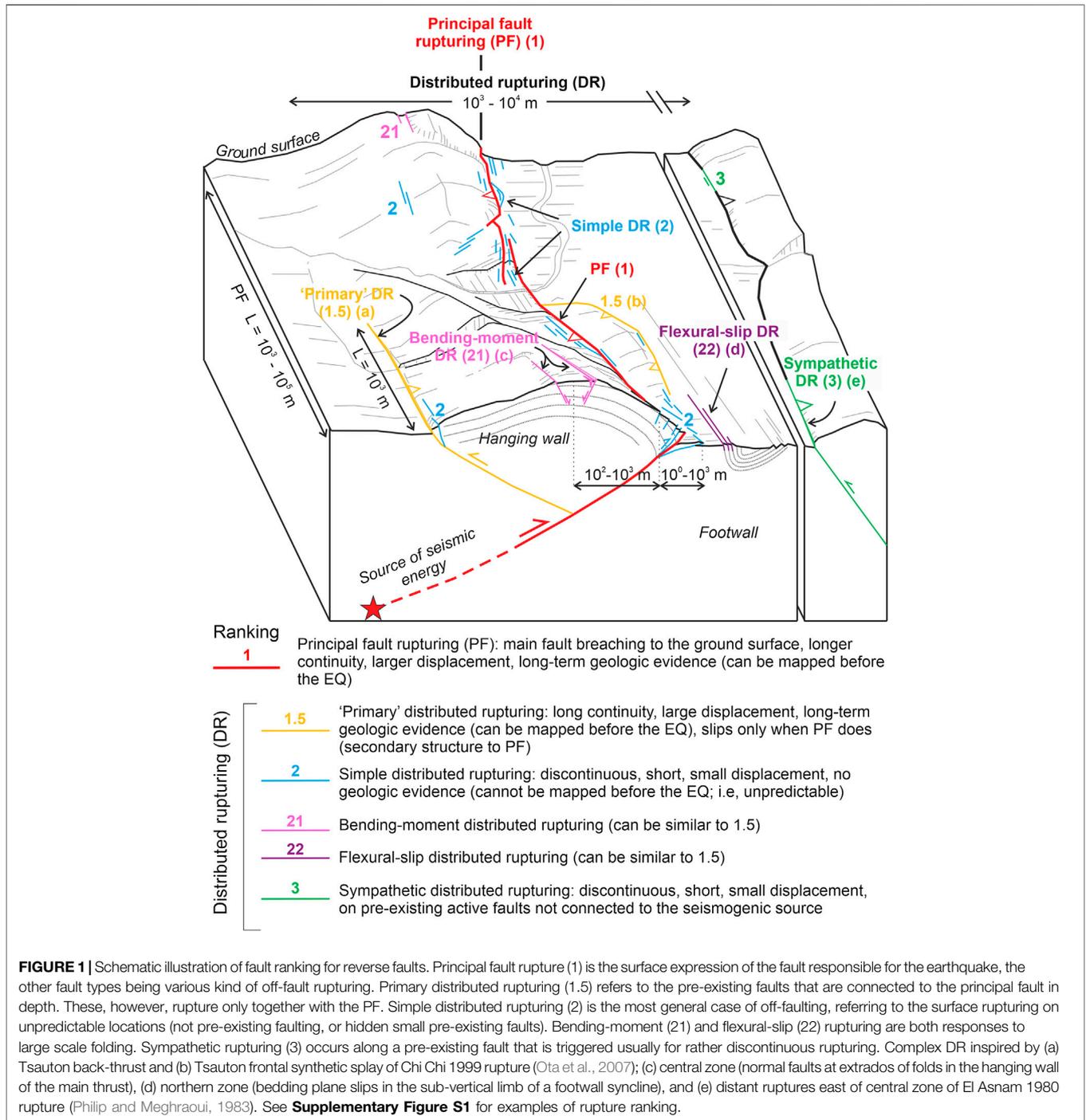
**Figure 2** shows the length distribution of the DR segments, and the number of DR segments per earthquake for hanging wall and footwall separately. When considering “all types” DR, the distribution of the rupture lengths in the hanging wall and footwall ranges from 1 to 3,665 m and 1 to 2,155 m, respectively, (**Figures 2A,B**); the mean length on hanging wall is 82.2 and 184.1 m on the footwall, and the 16th to 84th percentiles are 11.6–130.3 m on hanging wall, and 12.4–310.3 m on the footwall. Peaks in the simple DR length distributions are in the range of 2–25 m, the mean being 66.5 m for the hanging wall and 106.3 m for the footwall, and the 16th to 84th percentiles are 6.7 and 73.7 m for hanging wall, and 8.8 and 181.7 m for the footwall. The distribution of the number of DR

segments per earthquake when considering only the simple DR, in the hanging wall and footwall ranges, respectively, from 1 to 270 and from 1 to 41, the mean being 52 for the hanging wall and 9 for the footwall, and the 16th to 84th percentiles are 7 and 109 for hanging wall and 1 and 15 for the footwall (**Figures 2C,D**). The distributions in length and number of DRs and their spatial distribution are based on the maps published by the field working groups handled with the lowest level of interpretation possible.

The shapefiles of the surface ruptures and displacement data including the fault ranking of the 15 reverse earthquakes used in this work are included in the **Supplementary Table S1** and will be part of the next SURE database release. Data and results from this work contribute also to the Fault Displacement Hazard Initiative (FDHI) working group (Sarmiento et al., 2019; <https://www.risksciences.ucla.edu/nhr3/fdhi/home>).

## METHODOLOGY FOR PROBABILISTIC FAULT DISPLACEMENT HAZARD ANALYSIS

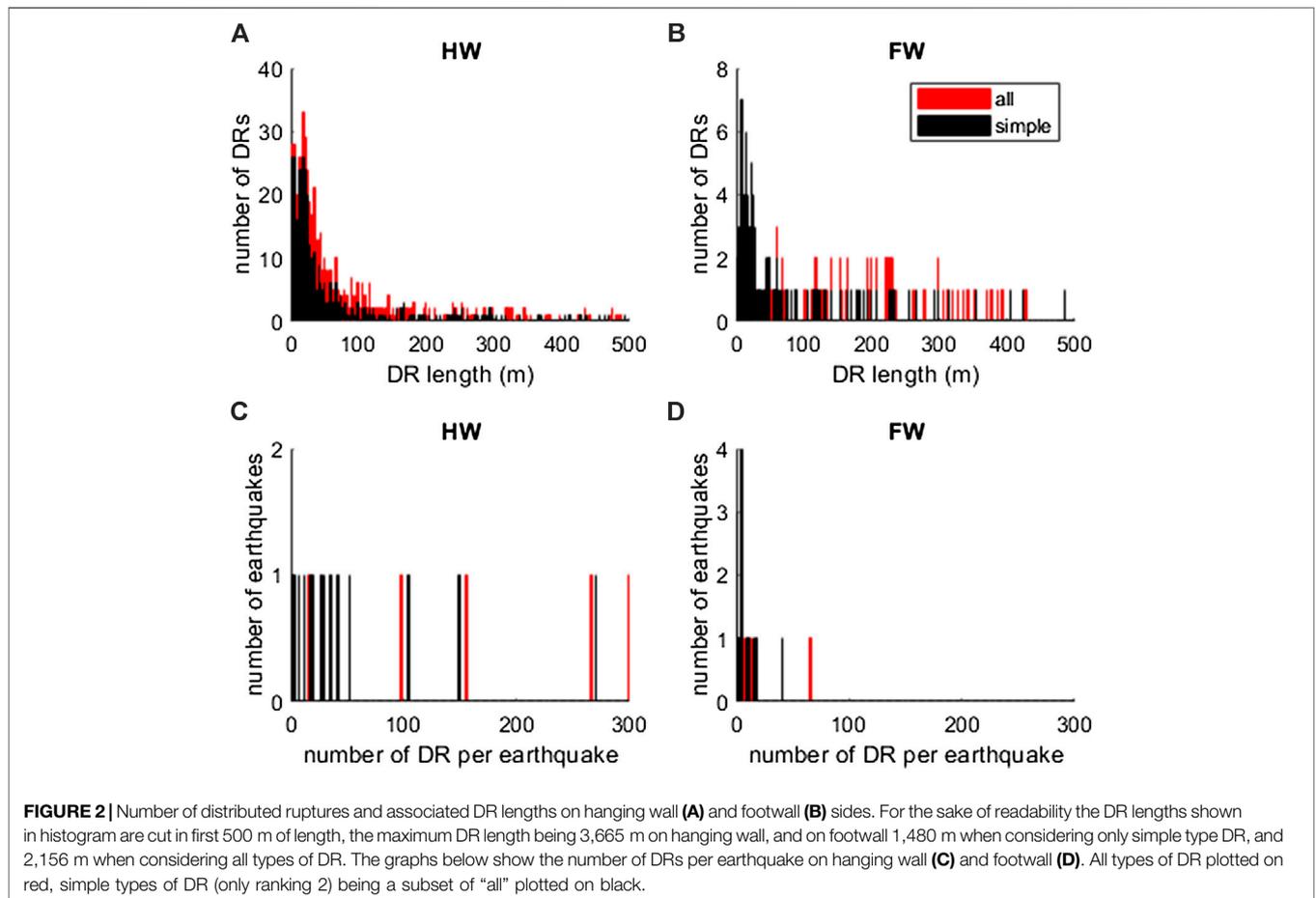
PFDDHA begins similarly to the traditional probabilistic seismic hazard assessment with seismic source identification, and



characterization of the earthquake frequency and distance distributions, but instead of ground motion distribution, continues with calculating the conditional probability of exceeding a displacement  $d$  (Youngs et al., 2003). Thus, for obtaining this conditional probability of exceeding a certain displacement level at a given distance from the PF, we need empirical attenuation relationships from the historical surface rupturing earthquakes describing the probability of DR occurrence and the displacement distribution as a function of

the distance from the PF. Previous works (e.g., Moss and Ross, 2011) have analyzed the PF faulting parameters for reverse faulting earthquakes, whereas in this work the emphasis is on obtaining the probability distribution and expected displacement of distributed rupturing. The methodology used for obtaining these parameters is described in the following.

In the approaches based on gridding (Youngs et al., 2003; Petersen et al., 2011), the probability of DR occurrence, also referred to as off-fault displacement to distinguish it from on-



fault rupturing for PF, is calculated for a given cell area located off the PF. The probability of a DR to pass through the site of interest is a function of the distance from the PF trace and of the surface area occupied by the site of interest. The grid approach to evaluate the probability of having a nonzero displacement within the area of interest is based on the rate of the observed number of cells with at least one rupture over the entire number of cells at the same distance from the PF. Logistic regression model was used by Youngs et al. (2003) to compute the conditional probability of distributed rupturing occurring at a point, using a functional form that incorporates the observed trends of decreasing frequency with increasing distance, increased density with increasing magnitude and lower frequency and faster decrease in frequency with increasing distance on the footwall than on the hanging wall side. According to the PFDHA methodology (Youngs et al., 2003; Petersen et al., 2011), the rate ( $\lambda$ ) at which the displacement ( $d$ ) on a site at a distance  $>0$  from a PF exceeds a specified level ( $x$ , with  $x > 0$ ) is conducted by solving the equation:

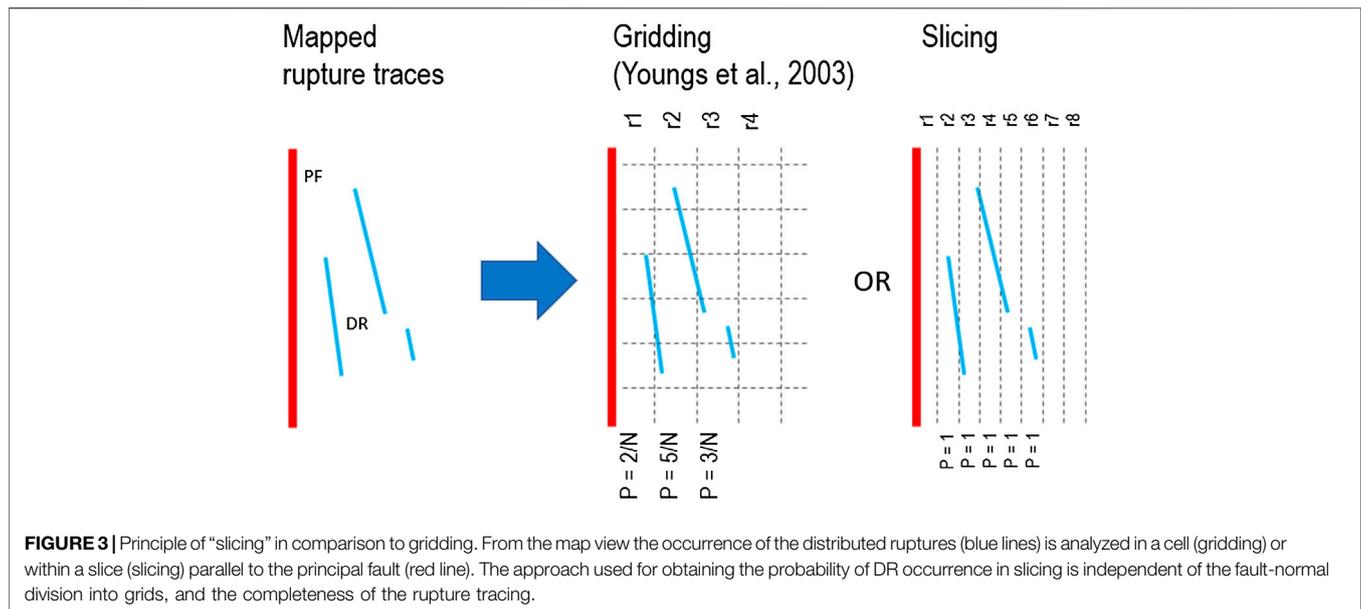
$$\begin{aligned} \lambda(d \geq x) &= \alpha(m) P[D > 0 | m] P[d > 0 | r, G] P[d > x | s, m, D_N], \end{aligned} \quad (1)$$

where

- (1)  $\alpha(m)$  is the annual rate of occurrence of an earthquake with magnitude  $m$ .
- (2)  $P [D > 0 | m]$  is the probability of having surface rupture on the PF, given that an earthquake with magnitude  $m$  occurred on the fault.
- (3)  $P [d > 0 | r, G]$  represents the probabilities of having a nonzero displacement at a distance  $r$  from the PF, as a function of a vector ( $G$ ) that include magnitude and that can additionally include the dimension of the site of interest and the expected total length of DRs.

In this paper we introduce the “slicing” method, more of which in the following chapter, in which the term  $P [d > 0 | r, G]$  is divided in two terms:  $P [d > 0 | r, m]$  and  $P [d > 0 | r, \text{dim}, F]$ . The former represents the probability of having a DR occurrence in a slice, being a function of the distance from the PF, with no regards of the dimensions of the site of interest. The latter expresses the conditional probability of having a DR within the site, given the occurrence of DR at the given distance. This term is a function of the dimensions of the site ( $\text{dim}$ ), and of the fraction ( $F$ ) of total DR length in respect to the length of the PF at the distance  $r$ , and

- (4)  $P [d > x | s, m, D_N]$  is the probability that the DR displacement exceeds a given value ( $x$ ), given  $r, m$  and  $D_N$ , where  $D_N$  is the displacement expected at a particular position



along the strike of the PF (to account for the variability in rupture location and displacement on the PF) and  $s$  represents the closest distance to that position from the DR.

For the purpose of this work, we assume that an earthquake with magnitude  $m$  produces surface faulting on the PF, and we focus on the third and fourth terms of Eq. 1 which require analyzing the spatial distribution of DR, deriving the most suitable measure of the displacement components along DR from the data collected in the database, estimating the probability of DR and finally estimating the probability of exceedance of a target level of displacement.

### Measurements From the Database: Distance Parameter $r$

We introduce here a new method called “slicing” for the analysis of the spatial distribution of DR. In the “slicing” approach, the area off the PF is divided in 10 m wide slices parallel to the PF strike. Slices containing at least a part of a DR segment are counted considering the distance from the PF (Figure 3). In practice, we first calculate the distance from DRs to the PFs by resampling the DRs and PF traces, filling in any gaps in coordinate data vectors greater than a defined 10 m tolerance apart in either dimension. We then calculate the minimum DRs to PF point-to-point distance. This point-to-point distance we name “ $r$ ” (Figure 4A). According to the ranking introduced in Figure 1, we calculate the distance  $r$  for a DR of

- (1) rank 2 with respect to rupture traces of rank 1 or 1.5, whichever the closest;
- (2) rank 1.5 with respect to rupture traces of rank 1; and
- (3) rank 21, 22, or 3 with respect to rupture traces of rank 1.

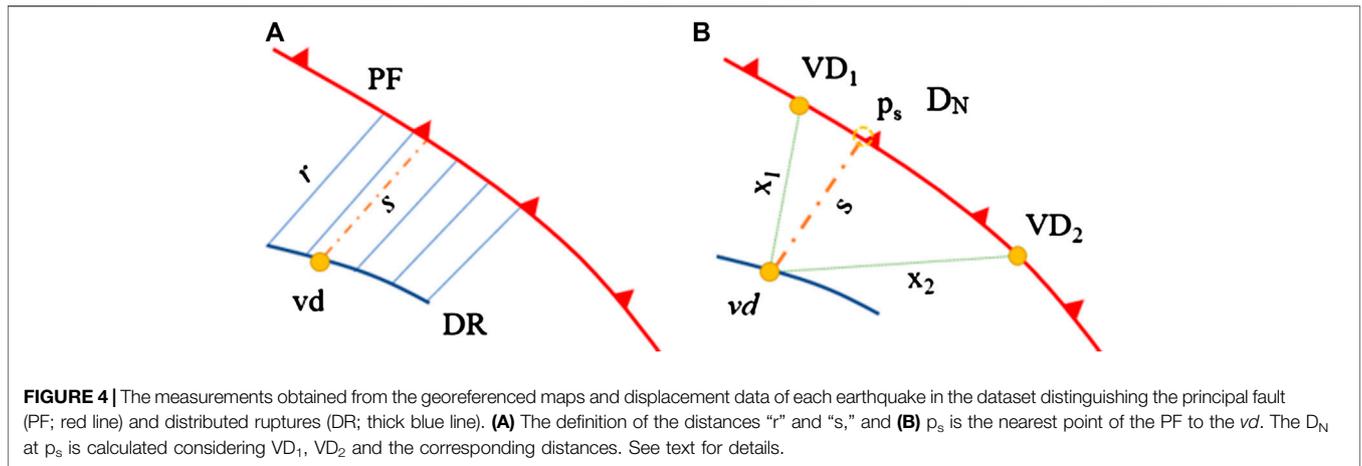
The width and interval of the slice, as well as the resampling tolerance are chosen, so as to reflect the resolution of the original maps used and still provide sufficient precision for the modeling.

The frequency–distance distributions of DRs is computed as the sum of slices intercepting at least a partial DR segment, normalized to the total number of the events. When the data of all earthquakes in the database are brought together, for each slice of distance  $r$  we can have a value ranging from 0 (none of the earthquakes has a rupture within the slice) to the total number of the events (all the earthquakes have at least a part of a DR segment intercepting the slice). This count is divided by the number of earthquakes to obtain the frequency. In Figure 5 we grouped the earthquakes according to the magnitude: 1)  $M_w$   $5 \leq M < M_w$  6 (including  $M_w$  4.9 Le Teil earthquake), 2)  $M_w$   $6 \leq M < M_w$  7, and 3)  $M \geq M_w$  7. Notice however, that the number of earthquakes per class of magnitude is small: 5, 6, and 4, respectively. The list of measurements  $r$ , their length, and the ranking of the DR as well as the ranking of the corresponding primary trace (1 or 1.5) are given in Supplementary Table S3.

Unlike the previous approaches utilizing gridding (Youngs et al., 2003; Petersen et al., 2011), the “slicing” method does not contain an assumption on the completeness of the database along the PF strike. We implicitly accept the very likely situation that not all the area is studied with the same precision in the field, as some parts can be hard if not impossible to reach. On the other hand, it is also likely that distributed rupturing does not occur homogeneously along the PF strike due to the physical factors, such as subterranean structures and material distribution, and the mechanics of the process. This will be accounted for when we introduce in the analysis an estimate of the relative coverage of the DR mapping with respect to the PF length, as will be shown later.

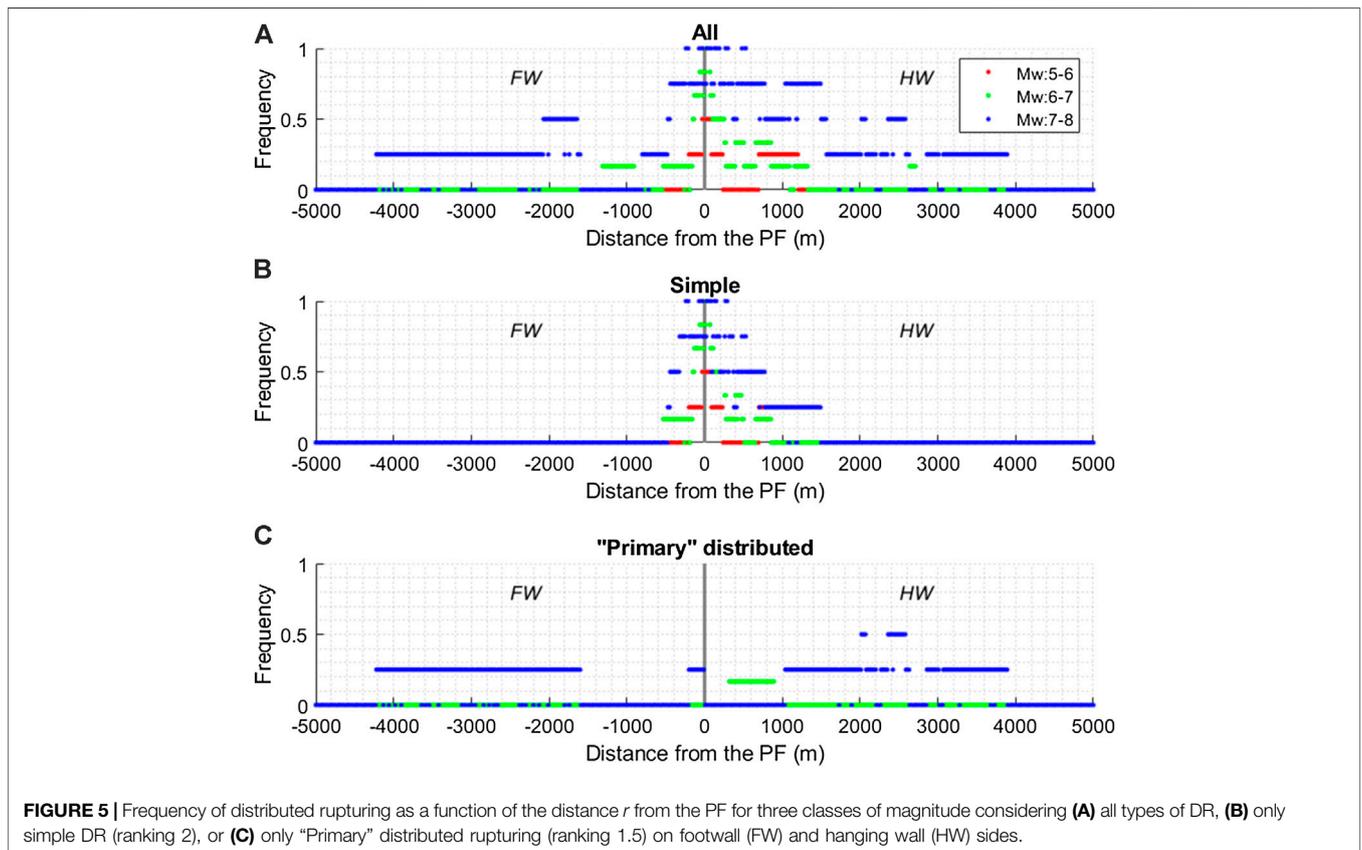
### Measurements From the Database: Vertical Displacement and the Distance Parameter $s$

*In situ* measurements are performed and reported in different ways from one study to another. The field conditions (visibility, accessibility, presence of displaced, and matching features) impact on the possibility to measure the



displacement, and the available slip components might not be uniform. In some, but not all cases, it is possible to derive all the slip components from available data. However, this is not always possible, and therefore some blanks remain in the database. For example, the net displacement ( $ND$ ), which represents the best the total deformation caused by the earthquake, can be calculated as a vector sum of the vertical displacement ( $VD$ ) and the two horizontal slip components (fault-parallel, and fault-normal slip), or obtained by utilizing

the fault dip angle if not all the vectors are known. Fault attitude (strike, slip) is usually well documented along the PF, but the DR can have different orientations with respect to the PF, and the assumptions of dip angles cannot be justified especially if not reported for the DR trace in consideration. Thus, obtaining rigorous  $ND$  based on reproducible mathematical methods is not equally correct when it comes to DR, as the DR slip parameters are rarely reported with a high level of detail.



Given that in the database the vertical slip component,  $VD$  for PF, and  $vd$  for DR, is the measure most often available for the surface ruptures of reverse faulting earthquakes, we decided to use this measure off displacement as a reference metric for the following analysis. This seems a reasonable choice as it is the most representative component of dip slip ruptures. Some distributed rupturing of the earthquakes considered in this work acted like normal faults near the surface, vertical slip measured in the field being opposite to the PF, but was considered here in absolute values. The database contains 234 measures of  $vd$ , most of them (53%) belonging to rank 2 DR, and 30% to 1.5 rank DR. The further analysis is limited to the “simple” DR (rank 2), to avoid biases due to folding-related or triggered DRs. Notice that the vertical slip values of the “primary” DR (rank 1.5) will be used in some cases later on for  $vd$  in the case where a rank 2 DR is closer to a rank 1.5 DR than to the PF.

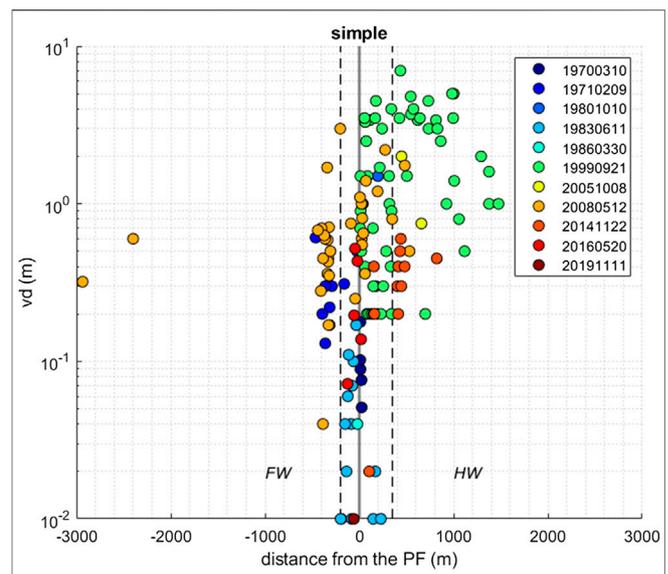
For the analysis of the spatial distribution of the vertical slip measured along the distributed surface ruptures, we utilize the nearest distance from the point of measurement to the corresponding primary trace, calling this distance  $s$  (Figure 4A). The difference of  $s$  in respect to  $r$  is that  $s$  is computed exactly and exclusively for the point on DR with an associated measure of vertical throw, whereas  $r$  is calculated by resampling the coordinates of the DR segments. For a sufficiently dense resampling procedure,  $s$  is a sub-sample of  $r$ .

The displacement measured or expected at a particular position along the strike of the PF ( $D_N$ ) could be obtained from slip profiling, but composing a profile is not a trivial procedure when faced with discontinuous fault traces and other geometrical complexities, such as overlapping branches and splays. Therefore, having the database of displacement points on both DR and PF, we propose an alternative method in which for each DR displacement point the parameter  $D_N$  is derived from the two nearest measured displacement points on PF (Figure 4B).  $D_N$  is defined by the two nearest vertical displacement values  $VD$  along the PF, inversely weighted by their distance to the point of  $vd$ .  $D_N$  is calculated for each point of vertical displacement along a DR as following:

$$D_N = (VD_1 x_2 + VD_2 x_1) / (x_1 + x_2), \quad (2)$$

where  $VD_1$  and  $VD_2$  are the nearest vertical displacement values along the corresponding primary trace (rank 1 or 1.5), and  $x_1$  and  $x_2$  are the corresponding distances from the point of  $vd$  to the points of  $VD$ . In some cases, the nearest point used to calculate the  $s$  ( $p_s$  in Figure 4B) is attributed to a segment different from the points of displacement along the PF used for calculating the  $D_N$ . In Supplementary Table S4, we list the coordinates of the  $VD_1$ ,  $VD_2$  and the end points of segments  $x_1$ ,  $x_2$ , and  $s$ .

In Figure 6, we show the  $vd$  of rank 2 DR with respect to their distance from the rank 1 and rank 1.5 ruptures. Figure 6 allows discussion of the spatial completeness of the  $vd$  with respect to the distance from the PF, in fact, for distances longer than few hundreds of meters both on the hanging wall and footwall, only large  $vd$  are mapped. A likely explanation is the difficulty of distinguishing DR of small dimension at large distances from the PF, and this must be kept in mind when using the empirical data.



**FIGURE 6 |** Analysis of the spatial distribution of the vertical displacement,  $vd$ , observations of simple DR (ranking 2) by earthquake in respect to the distance from the principal fault on hanging wall (HW) and on footwall (FW) sides. Dashed lines on FW and HW, in 200 and 350 m distance respectively, indicate the zone in which the database contains the largest variability of the values and can be assumed to be the most complete respect to the total DR occurred in reality.

## Probability of Distributed Rupturing

In this section we analyze the probability of DR first in a general situation, and then at a site with specific size in a defined distance  $r$  from the PF.

### Probability of a Distributed Rupturing Occurrence in a Slice

For the analysis of DR probability parameters (third term in Eq. 1), the distinction between PF and DR must be clearly defined. First of all, we establish a threshold distance within which all the ruptures are considered to be part of the PF. In this work the chosen distance is 5 m on both hanging wall and footwall sides, defining a 10 m wide PF zone. Utilizing this type of buffer zone seems reasonable also in respect to the observations in the field, where the surface material may not result in one well-defined, fully continuous principal fault trace. Indeed reverse faulting in areas with loose cover material induce often a broad flexural scarp, rather than a clear-cut scarp (e.g., Kelson et al., 2001; Yu et al., 2010). All the little and discontinuous ruptures within a distance inferior to 5 m to the PF were excluded from the further analysis of DR probabilities.

The “slicing” approach allows us to pose a simple question for each earthquake: “What is the probability to have at least one distributed surface rupture within a slice?” We then compute the occurrence of DRs in a slice for all the 15 earthquakes in the database, which means we are assuming the occurrence being independent from the position of the site along the PF strike. Subsequently, we apply a multinomial logistic model, which allows the analysis of various uncategorized variables, and that

**TABLE 2 |** Coefficients for Eqs 4 and 6.

Coefficient	Footwall	Hanging wall
<b>Equation 4</b>		
<i>a</i>	8.5431	2.9179
<i>b</i> <sub>1</sub>	-1.5586	-0.5566
<i>b</i> <sub>2</sub>	0.0099	0.0030
<b>Equation 6</b>		
<i>a</i>	-5.1043	-4.2549
<i>b</i> <sub>1</sub>	-0.6483	-0.1514
<i>c</i> <sub>1</sub>	0.1983	0.4404
<i>d</i> <sub>1</sub>	0.9461	0.5711
<i>std</i>	0.8812	0.9129

computes the relative risk of being in either the presence or absence of a DR from any given fault using a linear combination of predictor variables. Consequently, the probability of each outcome is expressed as a nonlinear function of *n* predictor variables:

$$\ln (P_j / (1 - P_j)) = a + b_1 X_1 + b_2 X_2 + b_n X_n, \quad (3)$$

where *P<sub>j</sub>* is the probability of an outcome being in category *j*, and *n* is the number of predictor variables. We focus on simple distributed rupturing, ranking category 2. The data we collected are organized into a two column matrix with the magnitude of the causative events and the classes of distance from the PF. The response is a vector of 1/0, with 1 where we have

at least a partial rank 2 DR and 0 where no rank 2 DR have been mapped. The global predictor matrix has 7,000 inputs, given by 15 earthquakes and 500 slices, which are centralized at distances from 5 to 5,005 m from the PF. In our case, Eq. 3 is

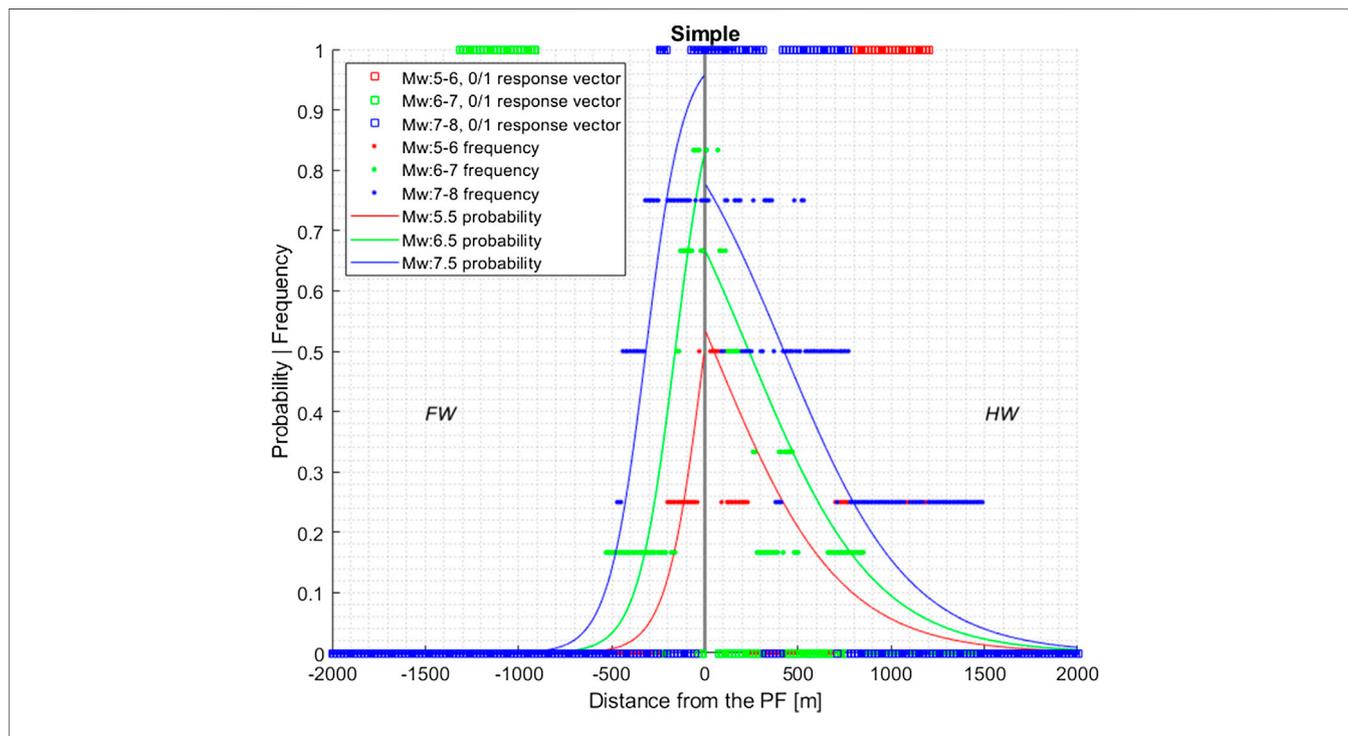
$$\ln (P_f / (1 - P_f)) = a + b_1 X_1 + b_2 X_2, \quad (4)$$

where *P<sub>f</sub>* is the probability of an outcome being in category “at least a partial rank 2 DR” with respect to the reference category of “no rank 2 DR,” and *X*<sub>1</sub> and *X*<sub>2</sub> are the earthquake magnitude and distance from the PF, respectively. The coefficients for Eq. 4 are given in Table 2.

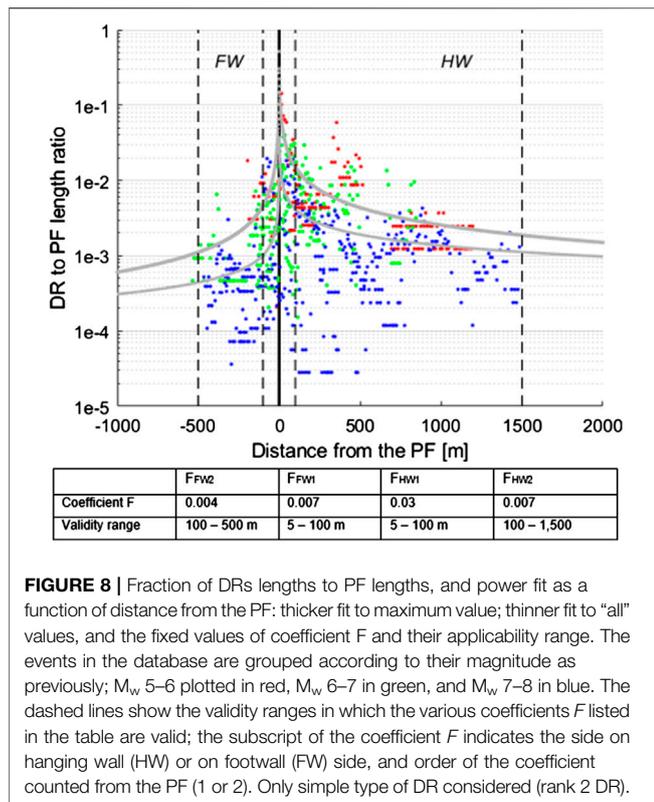
Figure 7 shows the predicted probabilities for the multinomial logistic regression model with predictors, *X*<sub>1</sub> and *X*<sub>2</sub>, and the coefficient estimates, *b*<sub>1</sub> and *b*<sub>2</sub>, for hanging wall and footwall. The probability curves are shown for three magnitude *M<sub>w</sub>* 5.5, *M<sub>w</sub>* 6.5, and *M<sub>w</sub>* 7.5. For a visual comparison, in Figure 6 we also show the frequency–distance distributions for the three classes of magnitude: *M<sub>w</sub>* 5 ≤ *M* < *M<sub>w</sub>* 6, *M<sub>w</sub>* 6 ≤ *M* < *M<sub>w</sub>* 7, and *M* ≥ *M<sub>w</sub>* 7. The number of the earthquakes is relatively small in each class of magnitude, but we can see clear differences between the classes due to which the division is considered reasonable anyway.

### Probability of Distributed Rupturing Occurrence at the Site

Distributed rupturing does not occur homogenously and continuously along the surface trace of the principal fault. To



**FIGURE 7 |** Logistic regression of the distributed rupturing in the database indicating the probability of observing at least a partial distributed rupture (DR) as a function of distance from a principal fault (PF) on hanging wall (HW) and footwall (FW) sides. The squares show the response to the logistic regression of rupturing either occurring (1) or not occurring (0) at a given distance. The points show the DR frequency in respect to the total number of events in each *M<sub>w</sub>* class, and the lines indicate the probability of at least a partial DR at a given distance.



estimate the probability of surface rupture occurrence at a site located at a distance from the PF we adopted a Monte Carlo approach. We discretized a space of length equal to the length of the PF in  $i$  sections of an arbitrary length of 1 m. Each section represents the possible center of a DR at a given distance off the PF. Next, we propose to calculate the number of secondary DR of given length considering the statistics of the database shown in **Figure 2**. Unfortunately, concerning the DR length statistics, the database does not allow to infer any relation between number and dimension of the surface ruptures relative to the parameters of the PF, such as magnitude or fault length.

Concerning the total length of DRs expected for a given PF length, we use the ratio observed in **Figure 8**. For each slice, for each earthquake, we computed the sum of the lengths of the DR traces, and normalized this sum to the length of the PF; this measure being the ratio *F*. The gray curves in **Figure 8** represent the fit of the maximum ratio at each distance (thick curve above) and the fit of the entire dataset (thin curve below). The rapid decay observed at a distance of  $\sim 100$  m from the PF both on hanging wall and footwall sides could be related to some physical reasons, but the bias in the data available does not allow to be confident in any particular extrapolation, and it could be, possibly, still related to the completeness of the database. For this reason, we consider a threshold distance of  $\sim 100$  m as a (soft) border between high to low ratio of DR in respect to the length of PF. Instead of considering either of the fit directly, we simplify the analysis by using two values of *F* at two ranges of distance from the PF based vaguely on the maximum ratio. For a distance  $< 100$  m, we can assume that the

expected length of the total DR will cover a fraction *F* of 0.03 of the length of the PF on hanging wall, and *F* of 0.007 on footwall. For a larger distance than 100 m, we assume *F* of 0.007 on hanging wall, and *F* of 0.004 on footwall. The range of applicability of these values in terms of distance from the PF is based on the reach of the database, being 500 m on footwall, and 1,500 m on hanging wall.

Given the total length of DRs and the range of individual length, we can derive the range of the expected number of DR. For example, for a 30 km long PF, at a 25 m-wide site located at a distance greater than 100 m from the PF, we expect the total length of the DR being 210 m ( $30 \text{ km} \times 0.007$ ). Assuming that the individual length of a DR ranges between the 16th and 84th percentiles of the observations, respectively of 7 and 74 m, we calculated the number of expected DR segments to be in the range of 3 ( $210/74$ ) to 30 ( $210/7$ ).

We then run  $10^4$  simulations, and for each simulation we

- extract the number ( $N$ ) of ruptures between the minimum-maximum number, by assuming a uniform distribution,
- deduce the length of each DR by dividing the total length of DR by  $N$ ,
- extract the  $N$  positions of the center of the surface ruptures from the  $i$  sampled centers,
- build the surface ruptures, by distributing the length of individual DR symmetrically in respect to their center, and
- verify if at least a rupture is located inside the site.

**Figure 9** illustrates an example of this Monte Carlo approach. In particular, **Figure 9** is a zoom of a 30 km fault for a site with a dimension parallel to the PF of 25 m, located at 500–525 m distance perpendicular to the PF. Note that the  $y$ -axis in **Figure 9** serves to stack the simulations (only the first 100 of the total of  $10^4$  are shown). The PF is shown as a red line on the bottom of the figure, whereas DRs are represented by the black lines parallel to the PF. To simplify the reading of the figure, the width of the site is represented by the blue lines, projected along the  $y$ -axis. The number of favourable cases (at least a partial rupture) is obtained by counting the number of at least one DR intersecting the site, i.e., how many times there is at least a partial black brick between the blue dotted lines. The total cases are the number of simulations.

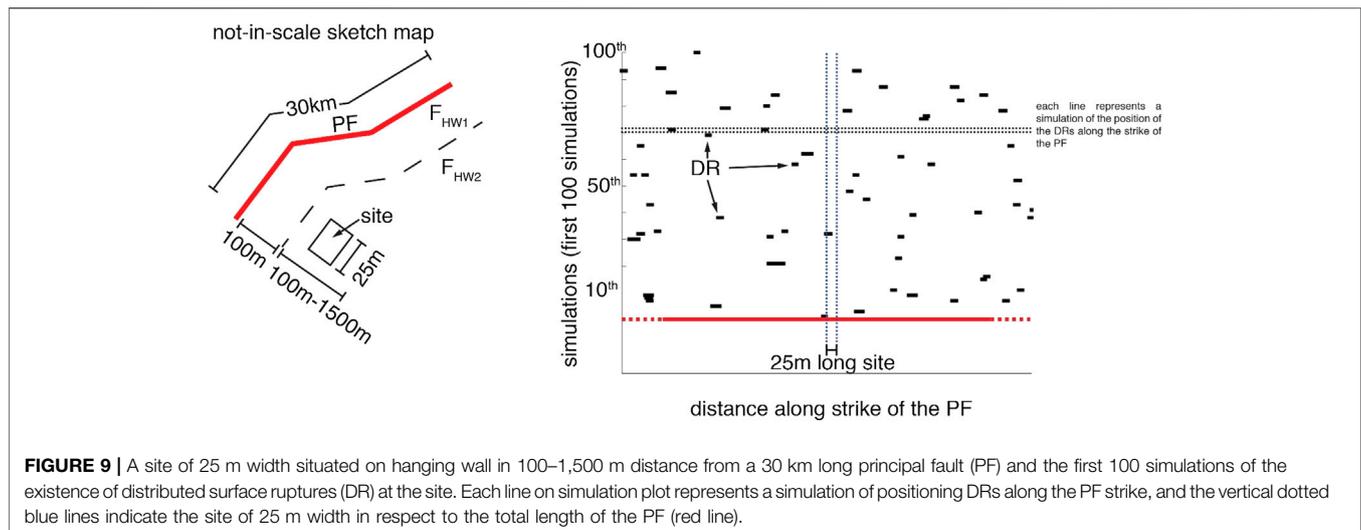
The conditional probability of DR occurrence within the site of interest  $P[d > 0 | r, \text{dim}, F]$ , that is a function of the distance  $r$  from the PF, of the dimension (dim) of the site, and of the length of the DR (that is the fraction *F* of the PF length), is then

$$P = \text{number of favourable cases} / \text{total cases} \quad (5)$$

Because of the sampling interval of the distance from the DRs to the PF (10 m), we assume that the probability *P* represents the probability for even a partial DR occurrence at a site with a dimension perpendicular to the PF of 10 m. This accuracy seems adequate to most engineering solutions the modeling may serve. The code is available in the **Supplementary Material** of this paper.

## Probability of Exceedance of Displacement Levels

The probability of exceedance of a level of displacement is obtained by analyzing the correlation between the amount of



displacement on DR vs. the distance from the PF, magnitude of the earthquakes and displacement on the PF. For this part, we use the distance parameter  $s$ , that indicates the distance from a point with a measured vertical surface slip along a DR to the nearest PF trace (**Figure 4B**). This distance is connected to the data point on DR in consideration.

Previous authors (Youngs et al., 2003; Petersen et al., 2011) chose the single value option to account for the magnitude of the event. Normalization to the maximum displacement ( $MD$ ) is a commonly used approach, and it is easily available value for every event. However, it does not consider the variability of the slip along the PF trace, which can be remarkable in large events with long surface rupture length (SRL) breaching varying surface conditions. An alternative option, adopted in this work, is to estimate the distributed displacement values by the amount of slip at the corresponding location along the PF, i.e., at (or very near) the other end of the  $s$  measurement line. This choice lets us correlate the off-fault displacement to the estimated local slip along the PF and associated slip patches in depth.

For the systematic analysis of the database, only the equal type of data could be confronted, thus the  $D_N$  should be derived from the similar slip parameters as the slip on DR considered. This approach considers only existing data; for the datapoints on DR with reported vertical slip parameter the two nearest points along PF with reported vertical displacement ( $VD$ ) are considered for calculating the parameter of expected displacement (not necessarily any two nearest points, if no  $VD$  available). This procedure allows us to account for the variability in rupture location and displacement along the PF.

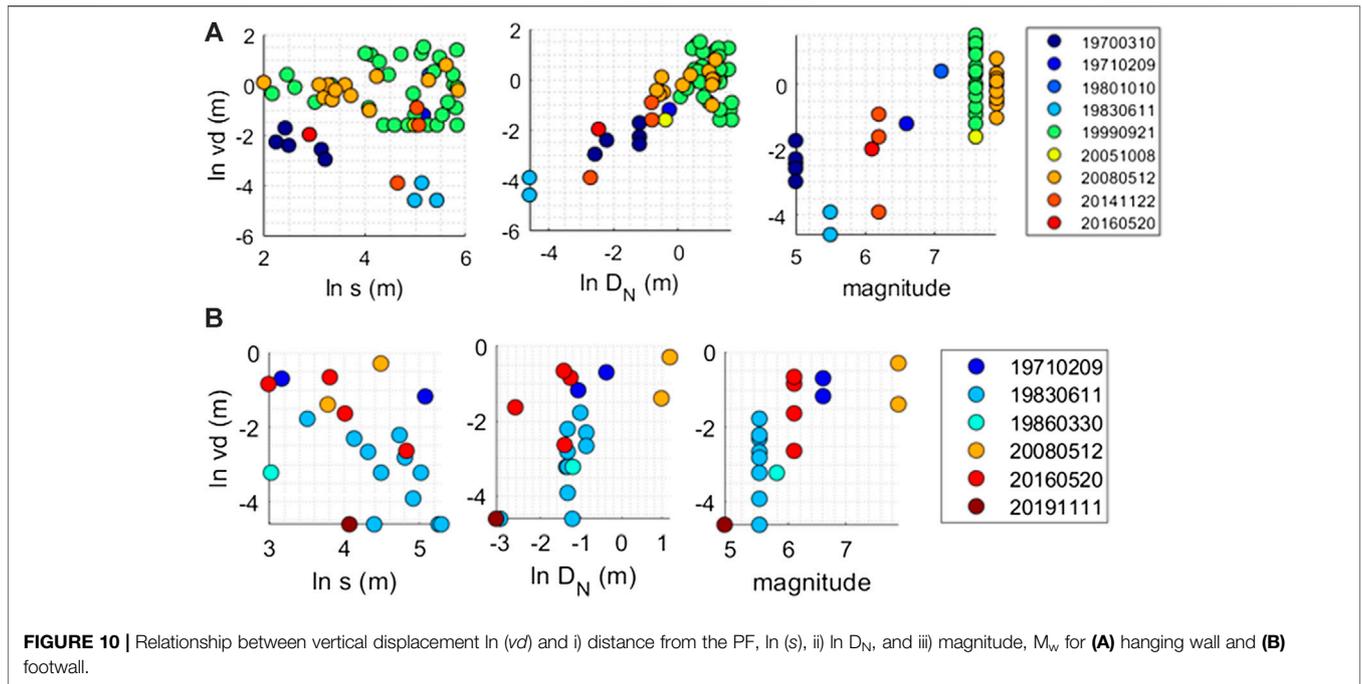
To evaluate the expected mean value of  $vd$  for rank 2 DR, we performed a regression analysis on  $vd$  and  $D_N$ , magnitude and distance  $s$  from the PF. Basing on the analysis of the spatial completeness of our database (**Figure 6**), we limited our analysis to a subset of data consisting of 48  $vd$  mapped at distances of 5–350 m on the hanging wall and of 21  $vd$  mapped at distances of 5–200 m on the footwall side, derived from 11 of the 15 earthquakes, with magnitudes from  $M_w$  4.9 to  $M_w$  7.9.

**Figures 10A,B** show the natural logarithm of  $vd$  in respect to the natural logarithm of distance  $s$  from the PF, the natural logarithm of  $D_N$  and magnitude, respectively for hanging wall and footwall. We can observe a weak correlation of  $vd$  with the distance both for hanging wall and footwall, and a linear correlation with  $D_N$  and magnitude. A possible reason for the weak correlation of  $vd$  with the distance from the PF is the short distances we are analyzing, and we do not expect a rapid decay in the first hundreds of meters. On the other hand, it is worth noting that we are correlating  $vd$  with an epicentral distance from the surface trace of the PF, and that we are ignoring possible important parameters as, for example, the dip of the fault, the distribution of the co-seismic slip at depth, the structural complexities of the principal faults (e.g., bends, gaps). All the distances in the database were measured from the georeferenced maps, in a 2D map view. Though this ignores the changes in the topography and the fault dip, potentially biasing the actual distances between the fault planes in steep slope regions, we do not expect excessive mistakes in fault displacement hazard analyses when performed equally on a map basis. The statistical processing of datasets is expected to smooth the potential errors by covering a wide range of geometrical arrangements between topography and fault dip.

The functional form of the equation used for the regression is

$$\ln(Y) = a + b_1(\ln(s)) + c_1(\ln(D_N)) + d_1(m), \quad (6)$$

where  $Y$  is the median expectation of  $vd$ ,  $a$  is the constant term,  $s$ ,  $D_N$ , and  $m$  are, respectively, the epicentral distance, the normalization factor derived from the vertical displacement of the PF, and moment magnitude. The values of the coefficients  $a$ ,  $b_1$ ,  $c_1$  and  $d_1$ , for hanging wall and footwall, and the residual standard deviations, are presented in **Table 2**. **Supplementary Figures S2A,B** show the residual plots for the regression of **Eq. 6** plotted to the distance from the PF. Even if the selected distances both on hanging wall and on footwall represent the reach we assume the empirical dataset is the most complete, there is quite large deviation in the vertical slip values.



We extrapolated the results from the analysis based on the near PF data also to distances greater than 200 and 350 m on footwall and hanging wall, respectively. Especially the large events may result  $vd$  greater than the model would suggest, and in general the empirical dataset contains more observations greater than the median, and greater the observations at short distances. This is somewhat counterintuitive but is likely due to the collection of the data in the fields, and it is possible that a certain distance from the PF only the DRs with largest vertical throw were observed and mapped.

Finally, in **Figure 11**, we show the probability of exceeding of vertical displacements at a site located at 500 m distance from the PF. We model the probability using a 3 sigma truncated distribution for two scenarios with  $M_w$  7 and  $M_w$  7.4, and with the maximum vertical throw expected on the PF, respectively of 2.3 and 3.85 m. The curves show the probability for each scenario for the site located either on hanging wall (solid lines), or on footwall (dashed lines).

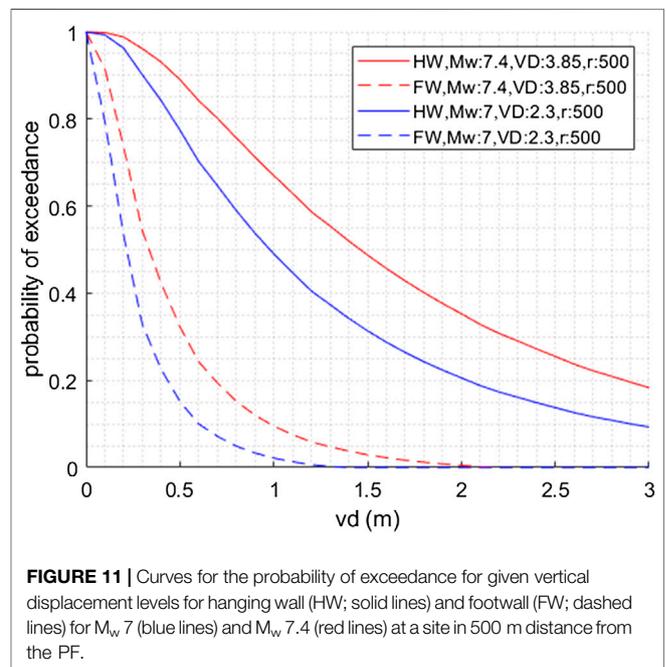
### CASE STUDY

In this work, we focused our attention on the development of new approaches to PFDHA by analyzing the hazard related to DR only, ignoring the probability of surface rupturing and slip distribution along PF. To appreciate the value of this new approach, we computed fault-displacement hazard curves for two scenario events, assuming a surface rupturing on PF along its total length. The probability of DR occurrence and displacement levels are estimated here for a site located on the hanging wall of our test case fault, at a distance from the PF greater than 100 m and for a characteristic earthquake scenario for which the  $M_w$  is derived from the total rupture length.

Matlab codes used for this case study are provided in the **Supplementary Material**.

### Suasselkä Post-Glacial Fault

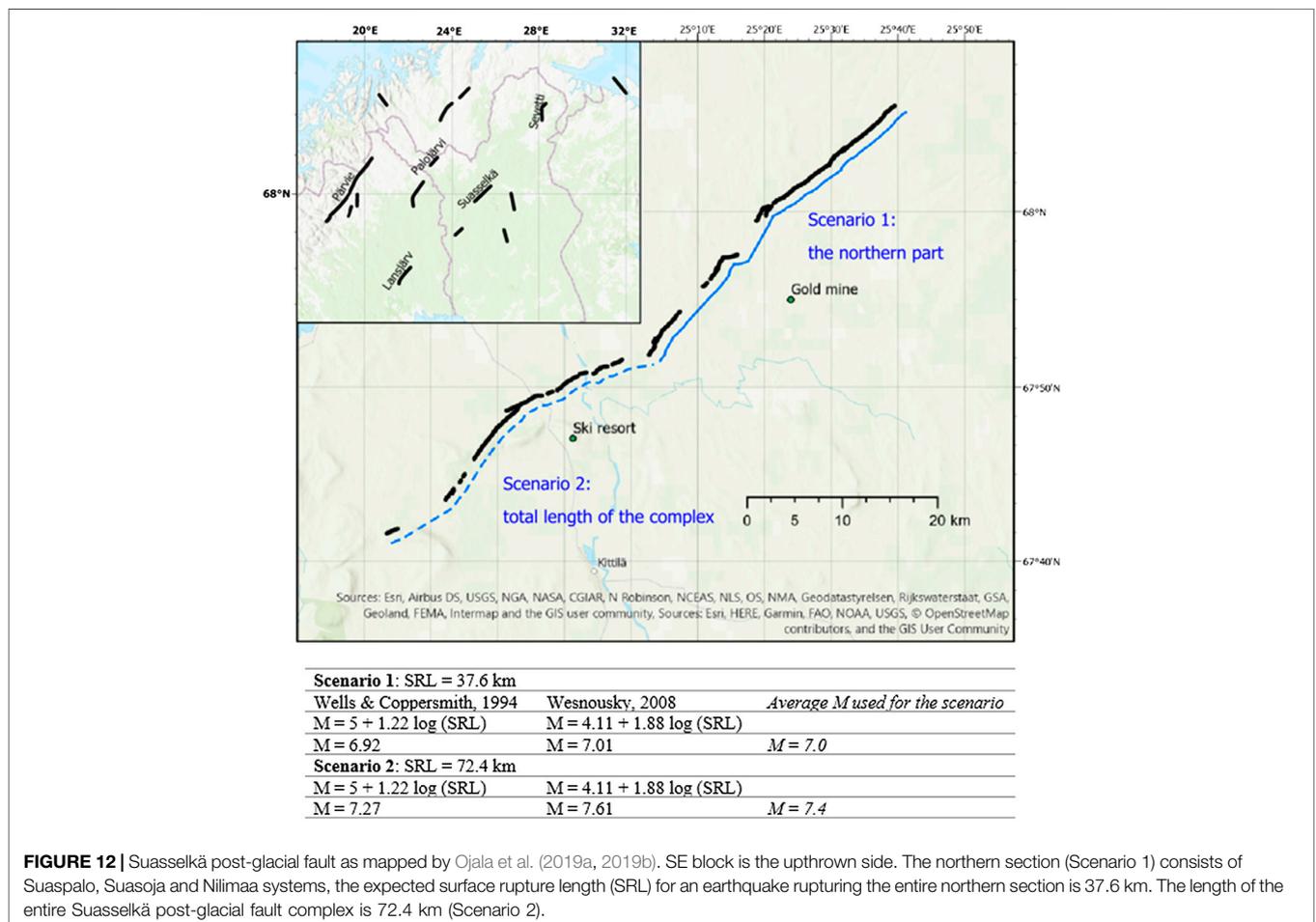
Northern Fennoscandia is characterized by large NE-SW orientated postglacial faults (PGF), which were generated in recent times (i.e., the last tens of thousands of years, during the Latest Pleistocene to Holocene) in response to NW-SE compressional stress (Lagerbäck and Sundh, 2008; Ojala et al., 2019b; Olesen et al., 2020). The PGFs of the area are mainly very



steep ( $45^{\circ}$ – $70^{\circ}$ ), and often associated with old pre-existing faults. According to Craig et al. (2016), the reverse-faulting earthquakes causing the large observed scarps were a response, like the coeval rapid uplift, to the Weichselian ice cap melting: those large events, with estimated magnitude from 7 up to 8 (see, e.g., Sutinen et al., 2014), happened at a time of extensional strain rate field. Surface geodetical measurements (i.e., GNSS) provide velocity field that strongly suggests current extension and rapid uplift (6.5 mm/year in the area of concern, according to the National Land Survey of Finland, 2018), whereas focal mechanisms show that compressional regime seems to be prevailing around the Scandinavian PGFs (Heidbach et al., 2018). Also the recent  $M_w$  4.6 shallow earthquake that hit the Kiruna mine area in Sweden (May 18, 2020), ~25 km east of the Pärvie PGF, shows a reverse focal mechanism as well on a steeply eastward dipping NNE-SSW striking fault.

Hence, we argue that the compressional environment that generated the PGF scarps is still active in the region. This is consistent with the recent findings of the Suasselkä fault complex, the longest PGF situated in Northern Finland. The drill core sampling of the fault proved that the ancient deformation zone has reactivated various times in the past, and the alteration of the Quaternary deposits indicates that the reactivation has continued

until recent times (Sutinen et al., 2014). Radiocarbon datings indicate three seismic episodes in the area, the most recent being 1–3 ky ago (Ojala et al., 2019b). In Suasselkä PGF complex, Ojala et al. (2019a) identified four fault systems composed of 37 isolated segments, segment lengths varying between 150 and 7,500 m, and estimated moment magnitudes ranging from  $M_w$  5.5 to 8.1 based on the rupture length and coseismic displacement considering the fault segments separately, and together as an entire complex. The nearest instrumentally observed earthquake to the Suasselkä fault complex took place in 2008 ( $M_w$  0.9) (Ojala et al., 2019b) and, in a general consideration, the recent strike-slip to reverse seismic activity has been very moderate in terms of magnitudes. To provide an insight on potential large earthquakes that may occur on this fault, we estimate their moment magnitude based on the PGF length and the scaling relationships. The most recent studies of Suasselkä PGF complex based on the LiDAR data and detailed field studies of the area (see Ojala et al., 2019a; Ojala et al., 2019b) result in SRL estimates longer than the previous assessments based mainly on visual inspection of the area (e.g., Paananen, 1987). The total SRL calculated between the fault complex end points is 72 km; doing so, we assume that the mapped segments are connected into a single crustal fault at depth.



## Example of Probabilistic Fault Displacement Hazard Analysis: Scenarios 1 and 2

In this work, the PFDHA modeling is applied to a hypothetical site of 10 m width  $\times$  20 m length (length = perpendicular to PF strike) situated on the hanging wall of the Suasselkä fault with the closest side at the distance of 500 m from the PF surface trace. The modeling is done for two possible scenarios: 1) rupturing of the northern half of the complex (i.e., Suaspalo, Suasoja, and Nilimaa faults), and 2) rupturing of the entire Suasselkä complex (Figure 12). We use the expected maximum vertical displacement of the scenario event on Suasselkä fault for transforming the attenuation parameters to the expected displacement values on DR, and thus the results are applicable in the section of the PF trace where the maximum displacement can be expected to occur. We are not making specific assumptions for the location of the maximum displacement along the PF strike. The magnitude is obtained taking an average of Wells and Coppersmith (1994) and Wesnousky (2008) equations for magnitude vs. SRL. The Moss and Ross (2011) empirical relationship is then used to calculate MD from magnitude, and then its vertical component ( $MD_v$ ) from fault dip. For scenario (1) the used  $M_w$  is 7.0, MD 2.37 m, and  $MD_v$  2.30 m (dip 75°; Ojala et al., 2019b). For scenario (2) the used parameters are  $M_w$  7.4, MD 3.99 m, and  $MD_v$  3.85 m.

### Scenario 1

The probabilities of having a rank 2 DR occurrence (Eq. 4) within the site of interest, extending from 500 to 520 m from the Suasselkä fault, are calculated at distances  $r_1$  and  $r_2$  from the PF of 505 and 515 m, respectively, on the hanging wall. The obtained probabilities are 0.408 ( $r_1 = 505$  m) and 0.401 ( $r_2 =$

515 m). The probability of having a DR occurrence in at least one of the two  $r$ , is  $1 - (1 - 0.408)(1 - 0.401)$ , and it is equal to 0.645. The probability of a rank 2 DR occurrence at the site is computed via Monte Carlo simulations (Scenario 2 section). It is independent from the distance  $r$ , and it is equal to 0.0182. Considering the size of the site under investigation (10 m width  $\times$  20 m length), the probability of DR occurrence is  $1 - (1 - 0.0182)^2$  and is equal to 0.0361. The conditional probability of exceedance is shown in Figure 13.

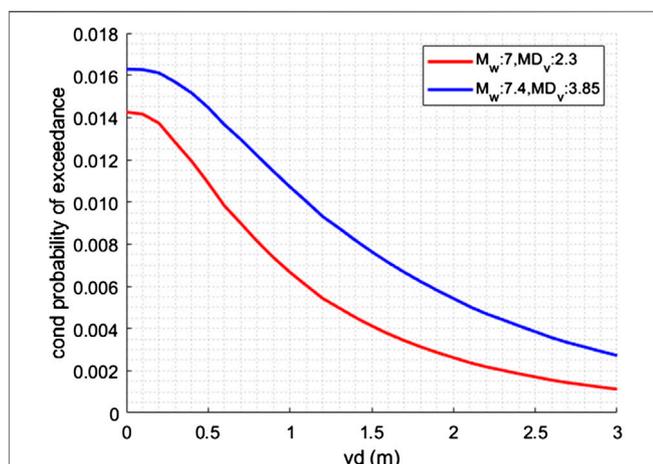
### Scenario 2

The probabilities of having a rank 2 DR occurrence (Eq. 4) on the hanging wall at distances  $r_1$  of 505 m and  $r_2$  of 515 m from the PF are, respectively, 0.496 and 0.489. The probability of having a DR occurrence in at least one of the two  $r$ , is  $1 - (1 - 0.496)(1 - 0.489)$ , and it is equal to  $\sim 0.7425$ . The probability of a rank 2 DR occurrence at the site, computed via Monte Carlo simulations, independent from the distance  $r$ , is equal to 0.02. For the site of interest (10 m width  $\times$  20 m length), the probability of DR occurrence at the site is  $1 - (1 - 0.02)^2$  and is equal to 0.0361. The conditional probability of exceedance is shown in Figure 13.

## DISCUSSION

The dataset used contains all the available surface rupture information of reverse faulting earthquakes published in the literature. The challenge regarding the homogenization of different types of the data in a single dataset is not new (Wells and Coppersmith, 1994). The first goal of such a study is to recognize the most accurate measure to consider for each earthquake. Displacement data used here covers 50 years of data collection and earthquakes in very different geological contexts, the type of data available varies from event to another, but also within an event depending on geologic and geographic conditions. The resolution of the maps and data provided by the field research groups have an impact on the resolution of this study. When working with datasets published by various study groups, the challenge lays also in bringing together different data formats and fault traces mapped at different scales. When it comes to the normalization parameter for displacement, common parameters of maximum displacement (MD) and average displacement (AD) along the PF are not reported in a homogenous way in the reference literature, especially the latter having more variety in its definition. The integral-based average considering the displacement distribution along the whole PF length would be the ideal factor describing the movement along the total PF rupture, but it is rarely provided and hardly obtained from the non-continuous data available. Here we propose to estimate  $D_N$  locally for each point in the database, which seems to be reasonably close to the slip profile-based approach but is remarkably faster to execute using the point-based database of surface displacement.

For a more reliable analysis, a standardized method for field work could diminish the uncertainty of further calculations and increase the accuracy of the future hazard analysis. Also, there is an evident need for harmonized nomenclature when it comes to



**FIGURE 13** | Conditional probability of exceeding a vertical displacement,  $vd$ , on a distributed rupture at a distance  $r = 510$  m from the principal fault (PF) on hanging wall side for two modeled scenarios on Suasselkä fault: scenario 1 of  $M_w$  7 and maximum vertical displacement on PF ( $MD_v$ ) of 2.3 m (red), and scenario 2 of  $M_w$  7.4 with a  $MD_v$  of 3.85 m (blue).

the fault displacement vectors and angles between them. Efforts have been made in the direction on harmonizing databases of surface rupturing data and the analysis done here is in accordance with the SURE structure (Baize et al., 2019), thus the parameters used in modeling can be updated as data is added into the database. Detailed field working of DR orientations and sub-surface structures will also help in more accurate ranking of the DR.

The variety of features observed in DR has led us to an in-depth analysis of their physical origin. This type of analysis has not been done in previous studies likely due to the relatively simpler nature of DR associated with strike-slip or normal faulting environments. The rank 2 DR in our database occur relatively close to the PF whereas rank >2 DR can occur at much greater distances. The occurrence of the more complex type of DRs (rank > 2) depends largely on the local structural geology, on the presence of pre-existing faults prone to be triggered (sympathetic faults), or large scale folding of hundreds of meters to kilometers in wavelength (possibly the origin of bending-moment or flexural-slip faulting). Following this analysis, we propose here to focus only on the rank 2 DR, which implies that our probability models can only be applied at short distances from PF. Extrapolation to distances greater than ~1 km from the PF should be considered with caution. Future studies will need to provide probability models that can account for rank >2 DRs for those sites that are potentially affected by bending moments or pre-existing structures.

Completeness issues on the spatial extent of the compile database have led us to propose an alternative approach to estimate the probability of DR at a given distance from the PF. The procedural enhancements introduced here in general aim at reducing the need of user interpretation. For example, for the DR distribution analysis by “slicing” is seen more independent from the field work, and it does not require analyzing the reason for the “blank” areas with no DRs on the surface rupturing maps.

A second issue regards the data collection procedure which becomes increasingly biased for distances greater than 200 and 350 m (footwall, hanging wall) from the PF. Indeed, extrapolating the displacement prediction equation to greater distances results in  $vd$  values lower than the values in the database. This is somewhat counterintuitive but is likely due to the collection of the data in the fields, and it is possible that a certain distance from the PF only the DRs with largest vertical throw were observed and mapped. Those, the equation we developed here are robust within few hundred meters from the PF, and extrapolation beyond these ranges need to be carefully considered.

To better understand the physical parameters controlling  $vd$ , it is worth noting the existence of several other factors in addition to the epicentral distance we are using here. For example, in the future one could introduce the dip of the fault and the distribution of the co-seismic slip at depth that can be found in the SRCMOD database (Mai and Thingbaijam, 2014), or the structural complexities of the principal faults (e.g., bends, gaps).

## CONCLUSIONS

Surface displacement data of 15 well-studied reverse earthquakes were gathered in an uniform database and used to derive empirical parameters used for describing the probabilities and expected levels of surface displacement in future events. The focus of this work was dedicated to enhancing the methodology for obtaining the parameters for distributed surface rupturing to be used in PFDHA on crustal reverse faulting tectonic regime.

Ranking of the distributed rupturing according to their complexity allows us to filter out the DR connected to the local structural setting in order to come up with parameters for modeling the simple, non-predictable DR that may take place anywhere along the strike of the PF, and on any reverse faulting earthquake where no structural complexities are expected.

In this work, we introduce the “slicing” methodology for analyzing the spatial distribution of DR to account for completeness issues of the empirical database. We divide the area off the PF in slices parallel to the PF strike and consider the slices containing at least a partial DR segment in order to generalize the probability and to avoid the bias occurring from the variability in the field mapping conditions that may cause either over or under sampling.

Vertical displacement data of various earthquakes are brought together obtaining a new approach to estimate the vertical displacement on the principal fault at the position nearest to the distributed rupture along which the vertical throw parameter is measured. The approach introduced here considers the alterations in the slip along the principal fault, and is derived directly from the point based database without slip profiling.

We combined empirical regressions and Monte Carlo simulations for calculating the probability of DR at the site at a chosen distance from the PF, and to estimate the probability of exceedance of the chosen vertical displacement level. To show the developed methodology, we provide Matlab codes applied to chosen scenario events along Suasselkä PGF.

## DATA AVAILABILITY STATEMENT

All datasets presented in this study are included in the article/**Supplementary Material**.

## AUTHOR CONTRIBUTIONS

FN compiled the database of reverse faulting earthquakes and carried out the implementation. PB and SB provided the geological expert opinion. FV, BP, AV, and OS performed the numerical simulations. All authors discussed the results and contributed to the final manuscript.

## SUPPLEMENTARY MATERIAL

The Supplementary Material for this article can be found online at: <https://www.frontiersin.org/articles/10.3389/feart.2020.581605/full#supplementary-material>

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**Conflict of Interest:** The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

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## GLOSSARY

**$D_N$**  expected vertical displacement calculated for each  $vd$  from the two nearest  $VD$

**DR** distributed (also called off-fault) rupturing: surface ruptures that are not direct surface expressions of PF faulting

**$F$**  fraction of total DR length in respect to the length of the PF of the corresponding earthquake

**FDHI** fault displacement hazard Initiative

**$ND$**  net displacement

**PF** principal (also called primary) fault rupturing; the fault along which the seismic energy releases in the seismic event

**PFDHA** probabilistic fault displacement hazard analysis

**PGF** post-glacial fault

**$p_s$**  the nearest point on the PF trace from the point of  $vd$  on DR to which the distance  $s$  is measured

**$r$**  a distance vector from DR trace to the nearest PF trace

**$s$**  a distance vector from a point along DR with displacement to the nearest PF trace

**SRL** surface rupture length

**SURE** a worldwide and unified database of surface ruptures

**$VD$**  vertical displacement on PF

**$vd$**  vertical displacement on DR

**VDL** vertical displacement level