Harmonization of real-world studies in multiple sclerosis: retrospective analysis from the RIReMS group

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ABSTRACT

Background. Worldwide multiple sclerosis (MS) centers have coordinated their efforts to use data acquired in clinical practice for real-world observational studies. In this retrospective study, we aim to harmonize outcome measures, and to evaluate their heterogeneity within the Rising Italian Researchers in MS (RIReMS) study group.

Methods. RIReMS members filled in a structured questionnaire evaluating the use of different outcome measures in clinical practice. Thereafter, thirty-four already-published papers from RIReMS centers were used for heterogeneity analyses, using the DerSimonian and Laird random-effects method to compute the between-study variance (τ^2).

Results. Based on questionnaire results, we defined basic modules for diagnosis and follow-up, consisting of outcome measures recorded by all participating centers at the time of diagnosis, and, then, at least annually; we also defined more detailed/optional modules, with outcome measures recorded less frequently and/or in the presence of specific clinical indications. Looking at heterogeneity, we found 5-year variance in age at onset (ES=27.34; 95%CI=26.18, 28.49; p<0.01; τ^2 =4.76), and 7% in female percent (ES=66.42; 95%CI=63.08, 69.76; p<0.01; τ^2 =7.15). EDSS variance was 0.2 in studies including patients with average age <36.1 years (ES=1.96; 95%CI=1.69, 2.24; p<0.01; τ^2 =0.19), or from 36.8 to 41.1 years (ES=2.70; 95%CI=2.39, 3.01; p<0.01; τ^2 =0.18), but increased to 3 in studies including patients aged >41.4 years (ES=4.37; 95%CI=3.40, 5.35; p<0.01; τ^2 =2.96). The lowest variance of relapse rate was found in studies with follow-up duration <2 years

(ES=9.07; 95%CI=5.21, 12.93; p=0.02; τ^2 =5.53), whilst the lowest variance in EDSS progression was found in studies with follow-up duration >2 years (ES=5.41; 95%CI=3.22, 7.60; p=0.02; τ^2 =1.00). **Discussion**. We suggest common sets of biomarkers to be acquired in clinical practice, that can be used for research purposes. Also, we provide researchers with specific indications for improving inclusion criteria and data analysis, ultimately allowing data harmonization and high-quality collaborative studies.

KEY WORDS: multiple sclerosis; real world; harmonization; outcome measures.

INTRODUCTION

Multiple sclerosis (MS) is a chronic immuno-mediated disease of the central nervous system, characterised by young age at onset, female predominance, and heterogeneous symptoms (Thompson et al., 2018). As such, enormous effort has been put into the identification and characterisation of the variety of MS symptoms (e.g., motor, cognitive, sensory, etc) through disease progression, using clinical, cognitive, neuroimaging, laboratory, and neurophysiological outcome measures (Tur et al., 2018).

To overcome sample size constraints, worldwide MS centres have coordinated their efforts to use data acquired in clinical practice for real-world observational studies, mostly aiming at identifying predictors of poor outcome and treatment response/failure (Glaser et al., 2019). However, in the absence of a consensus, real-world studies generally include the limited number of variables recorded in clinical practice by all participating centres (Middleton et al., 2018; Trojano et al., 2017). As such, the harmonization of outcome measures collected in clinical practice could be helpful to produce more granular real-world studies. In a previous study, Salter and colleagues showed the feasibility of harmonization between North American Research Committee on MS (NARCOMS) Registry, German MS Register (GMSR), and United Kingdom MS (UK-MS) Register (Salter et al., 2020). However, MS registries and real-world studies are composed by data collected from different MS centres, that are not necessarily homogeneous (e.g., selection bias) (Fortier et al., 2017; Salter et al., 2020; Trojano et al., 2017).

Among collaborative research networks contributing to real world studies, the Rising Italian Researchers in MS (RIReMS) (www.rirems.it) study group is composed of experienced neurologists from representative Italian MS Centres, and has already promoted many cross-sectional and

longitudinal real-world studies (Ferraro et al., 2020a; Lanzillo et al., 2018). Thus, studying the extent to which populations across RIReMS sites are homogeneous could be helpful for future study design within RIReMS and other collaborative networks. In this retrospective study, we aim to: (1) evaluate preferred clinical, cognitive, neuroimaging, laboratory, and neurophysiological outcome measures in RIReMS centres; (2) define modular sets of biomarkers for the clinical practice and for research purposes in RIReMS centres; and (3) measure the heterogeneity of MS populations included in different studies from RIReMS centres.

METHODS

Study design

The present study was conducted within 20 RIReMS centres. In the first part of the study, RIReMS members were asked to fill in a structured questionnaire evaluating outcome measures used in clinical practice and/or for research purposes at their MS Centres, which were used to define modular sets of outcome measures. Then, clinical studies published by RIReMS centres were used for retrospective heterogeneity analyses. Results were discussed during a RIReMS meeting in January 2020 in advance of manuscript preparation.

Questionnaire for outcome measures

To define clinical, cognitive, neuroimaging, laboratory, and neurophysiological outcome measures (aim 1), each RIReMS member was required to fill in a structured questionnaire (English version in **Supplementary Material 1**). This questionnaire consisted of a list of demographics, clinical features (e.g., examination, scales, etc), cognitive tests, patient-reported outcome measures (PROMs), laboratory analyses, neurophysiological exams, and MRI parameters. For each outcome measure, participants were required to report: (1) whether they had been using them for their clinical practice and/or research purposes; (2) when and how frequently (e.g., diagnosis, follow-up, annually, in the presence of specific symptoms); (3) and their usefulness. Participants were also asked to share any additional thoughts in an open question.

Modular sets of outcome measures

Based on this questionnaire, modular sets of assessments were defined (aim 2). In particular, basic modules for diagnosis and for follow-up consisted of outcome measures recorded by all participating centres at the time of diagnosis, and, then, at least annually during follow-up. The

detailed module included outcome measures recorded in most participating centres in the presence of specific clinical indications. Finally, we also defined an optional extension module, including outcome measures recorded in some participating centres in the presence of specific clinical scenarios.

Heterogeneity analyses

Each participating RIReMS member was invited to suggest up to three previously-published papers with the following characteristics: 1) independent populations; 2) publication in the past 10 years; 3) inclusion of a wide range of outcome measures. Overall, participating RIReMS members suggested thirty-four papers, which were reviewed for consistency with inclusion criteria by two independent assessors and, then, were included in the analyses (Table 1) (Bajrami et al., 2018; Bisecco et al., 2018; Buscarinu et al., 2017; Calabrese et al., 2017; Camerota et al., 2017; Carotenuto et al., 2019; Cocozza et al., 2019; Coghe et al., 2018; De Biasi et al., 2016; Della Corte et al., 2018; Di Filippo et al., 2014; Di Gregorio et al., 2018; Ferraro et al., 2020b; Gaetani et al., 2019; Gajofatto et al., 2014; Lanzillo et al., 2017; Lipp et al., 2020; Lorefice et al., 2019; Mallucci et al., 2018; Malucchi et al., 2017; Manni et al., 2019; Marastoni et al., 2019; Megna et al., 2019; Moccia et al., 2019, 2016; Paolicelli et al., 2016; Pellegrino et al., 2018; Realmuto et al., 2019; Salemi et al., 2019; Scalfari et al., 2018; Sola et al., 2010; Vercellino et al., 2009; Zoccolella et al., 2012). From these studies, we extracted mean (and standard deviation), and rates of different outcome measures, as appropriate. Average age at onset, if not directly available, was calculated as the difference between age and disease duration reported in the study; standard deviation of the age was then included. If data was presented differently (e.g., median and range), conversion was performed with previously described methods (Wan et al., 2014). If the selected study was multicentre, data extraction and conversion only covered the sub-population from the RIReMS centre.

We preliminary summarised demographics and clinical features from different studies and, since no obvious differences were detected (e.g., outliers), all studies were included in subsequent analyses. To retrospectively evaluate heterogeneity in demographics and clinical features (aim 3), we estimated the between-study variance (τ^2) using the DerSimonian and Laird random-effects method, which includes studies' effect estimates and standard errors as input, as in a previous similar study (Salter et al., 2020); weights from these analyses were used for forest plot graphical presentation. We also estimated the percentage of between-study heterogeneity that is attributable to variability in the true treatment effect (I^2). As such, we obtained the amount of heterogeneity (τ^2), and the percent this heterogeneity was actually true, and not related to sampling variation (I^2) (Harris et al., 2008; Salter et al., 2020). These statistical models were applied to demographics and clinical features, and also to sub-analyses by subgroups (e.g., age, study duration). Results are presented as effect size (ES), 95% confidence intervals (95%CI), and p-value. Statistical significance was set at p<0.05. Statistical analyses were performed using Stata 15.0.

RESULTS

Outcome measures and modular sets of biomarkers

Based on the questionnaire for the use of different outcome measures, we defined four modular sets: basic module at diagnosis, basic module at follow-up, detailed module, and optional extension module. Full details of modular sets are reported in **Figure 1**.

Heterogeneity of RIReMS studies

Looking at demographics on RRMS patients, we found a variance of 5 years in age at onset (n=16) (ES=27.34; 95%CI=26.18, 28.49; p<0.01; τ^2 =4.76; I²=91.70%) (**Figure 2a**). We found a variance of 7% in the percent of females (n=18) (ES=66.42; 95%CI=63.08, 69.76; p<0.01; τ^2 =7.15), only in part attributable to actual heterogeneity (I²=51.70%) (**Figure 2b**).

Looking at motor disability on the whole population (n=34), we found a variance of 2.7 in expanded disability status scale (EDSS) (ES=3.08; 95%Cl=2.50, 3.65; p<0.01; τ^2 =2.73; l²=99.70%), which decreased to 0.2 when considering studies including patients with average age <36.1 years (n=11) (ES=1.96; 95%Cl=1.69, 2.24; p<0.01; τ^2 =0.19; l²=96.30%), or from 36.8 to 41.1 years (n=9) (ES=2.70; 95%Cl=2.39, 3.01; p<0.01; τ^2 =0.18; l²=90.12%); on the contrary, EDSS variance increased to 3, when considering studies on more advanced disease stages (age>41.4 years) (n=12) (ES=4.37; 95%Cl=3.40, 5.35; p<0.01; τ^2 =2.96; l²=99.70%) (**Figure 3**). Looking at more specific measures of motor disability, we found a variance of 41 seconds in time 25-feet walking test (n=5) (ES=15.19; 95%Cl=9.31, 21.06; p<0.01; τ^2 =41.61; l²=97.21%), and of 27 seconds in 9-hole peg test (n=3) (ES=30.70; 95%Cl=23.72, 37.68; p<0.01; τ^2 =27.24; l²=83.40%).

Looking at longitudinal studies, we found a variance of 19.4% in the yearly rate of patients with relapses (n=8) (ES=10.74; 95%CI=6.96, 14.53; p<0.01; τ^2 =19.47; I²=70.11%), which decreased to 5.5% in studies with follow-up duration <2 years (n=5) (ES=9.07; 95%CI=5.21, 12.93; p=0.02; τ^2 =5.53; I²=28.54%), and increased to 31.7% in studies with follow-up duration >2 years (n=3) (ES=12.26; 95%CI=5.41, 19.11; p<0.01; τ^2 =31.72; I²=86.65%) (**Figure 4a**). On the contrary, we found a variance of 5.8% in the yearly rate of patients with EDSS progression (n=6) (ES=4.81; 95%CI=2.21, 7.42; p=0.03; τ^2 =5.83; I²=59.07%), which increased to 15.6% in studies with follow-up duration ≤2 years (n=3) (ES=4.59; 95%CI=-0.74, 9.92; p=0.02; τ^2 =15.68; I²=71.90%), and decreased to 1% in studies with follow-up duration >2 years (n=3) (ES=4.59; 95%CI=-0.74, 9.92; p=0.02; τ^2 =15.68; I²=71.90%), and decreased to 1% in studies with follow-up duration >2 years (n=3) (ES=4.59; 95%CI=-0.74, 9.92; p=0.02; τ^2 =15.68; I²=71.90%), and decreased to 1% in studies with follow-up duration >2 years (n=3) (ES=4.59; 95%CI=-0.74, 9.92; p=0.02; τ^2 =15.68; I²=71.90%), and decreased to 1% in studies with follow-up duration >2 years (n=3) (ES=4.59; 95%CI=-0.74, 9.92; p=0.02; τ^2 =15.68; I²=71.90%), and decreased to 1% in studies with follow-up duration >2 years (n=3) (ES=5.41; 95%CI=3.22, 7.60; p=0.02; τ^2 =1.00; I²=0.00%) (**Figure 4b**).

Looking at MRI measures, we found a variance of 402 mL in brain volume (n=3) (ES=1354.30; 95%CI=1142.35, 1566.26; p=0.09; τ^2 =402.67; I²=57.60%), and 245 mL in grey matter volume (n=4) (ES=727.36; 95%CI=636.71, 818.00; p<0.01; τ^2 =245.44; I²=99.60%).

Looking at cognitive tests, we found a variance of 17 in symbol digit modalities test score (n=7) (ES=40.34; 95%CI=34.24, 46.44; p<0.01; τ^2 =17.67; I²=96.35%), and of 25 in paced auditory serial addition test (3 seconds) score (n=6) (ES=34.72; 95%CI=30.46, 38.98; p<0.01; τ^2 =25.41; I²=94.14%).

Looking at PROMs, we found a variance of 18 in Beck depression inventory (n=4) (ES=11.56; 95%CI=7.16, 15.96; p<0.01; τ^2 =18.8; I²=94.05%), of 39 in fatigue severity scale (n=4) (ES=40.59; 95%CI=31.39, 49.80; p<0.01; τ^2 =39.70; I²=95.10%), and of 6.5 in modified fatigue impact scale (n=4) (ES=38.84; 95%CI=33.57, 44.10; p=0.96; τ^2 =0.0; I²=0.0%).

DISCUSSION

In the present retrospective study, we compared the use of clinical, cognitive, neuroimaging, laboratory, and neurophysiological outcome measures between 20 RIReMS centres, and defined modular sets of biomarkers for the clinical practice and for research purposes. Also, we evaluated the variance of different outcome measures between RIReMS centres, and showed that these populations can be combined reliably, but with some caveats. As such, present results will be helpful for future study design, and will possibly improve harmonization of data collection in clinical practice.

We showed that all RIReMS centres have collected a number of basic outcome measures at the time of diagnosis, and, then, over the follow-up (at least annually), from patients seen routinely in outpatient clinics, irrespectively of whether they were currently recruited for a study. Our basic modules are in line with the current standard of clinical registries (Glaser et al., 2019; Pugliatti et al., 2012). Also, based on the experience of RIReMS centres, we developed detailed and optional extension modules with suggestions for using specific tools/scales to evaluate MS symptoms. Looking at outcome measures selected for different modules, we found a consensus about the functions/domains that should be assessed, but, less so, about the nature of specific assessments. This could be explained by differences in study designs, preferences based on previous experience, and/or availability of resources. However, our study also highlights areas of improvements for assessing specific symptoms. For instance, MS-specific scales for pain and sexual dysfunction are not validated yet in Italian language. Also, the collection of some variables (e.g., comorbidities) could be improved with specific classifications. Not least, outcome measures from detailed and optional extension modules might be biased by the collection depending on the clinical indications (e.g., ophthalmological evaluations in patients treated with fingolimod, spasticity scales in patients

treated with cannabinoids), and, thus, should be interpreted cautiously. Alternatively, basic modules could be expanded, with the inclusion of a wider range of outcome measures to be collected in clinical practice and, then, used in real-world studies. Of course, the number of participants, and the available staff (and costs) should be considered when extending the assessments to detailed and optional extension modules in clinical practice.

Thirty-four previously-published studies from RIReMS centres overall showed good homogeneity, but with some caveats. Age at onset presented with 5-year variance, mostly coming from the inclusion of studies conducted in different years; in particular, older age at onset was found in studies conducted in less recent years (Vercellino et al., 2009), or including historical cohorts (Moccia et al., 2016), when compared with contemporary populations, where earlier diagnosis is made possible by the application of newer diagnostic criteria (Brownlee et al., 2015; Petruzzo et al., 2020). On the contrary, the 7% variance of females is related to actual sampling variations, possibly as a consequence of studies with smaller sample size (Malucchi et al., 2017). Disability measured with the EDSS is known to be highly dependent on age (Manouchehrinia et al., 2017), and, accordingly, we found very small variance when including selected age ranges. The largest variability of disability was found in more advanced disease stages, when individual trajectories of progression have diverged (Bodini et al., 2020; Signori et al., 2018). Looking at longitudinal measures, variance was smaller within 2 years from study inclusion for relapses, and after 2 years for EDSS progression. These findings are in line with previous studies suggesting that, over time, the number of relapses progressively reduce (increasing population variability) (Schwehr et al., 2018), whilst disability outcomes need time to fully disclose (Kalincik et al., 2015). Of note, we did not perform a systematic review and meta-analysis of previously-published papers from RIReMS members (e.g., PRISMA checklist and flow diagram), which was out of the scope of the present manuscript, but relied on

participating members' suggestions for paper selection to depict heterogeneity of demographics and MS clinical features in different RIReMS research populations.

Based on the lesson learned from our previous experiences and from our retrospective heterogeneity analyses, we have integrated MS-specific recommendations to the Maelstrom Research guidelines for rigorous retrospective data harmonization (Fortier et al., 2017). The Maelstrom guidelines consist of a six-step checklist to achieve a successful harmonization, from dataset generation to result dissemination (Fortier et al., 2017). Here, we have developed MSspecific suggestions for each step of Maelstrom guidelines (**Figure 5**):

- Step 0: define the questions and objectives. Research questions and study objectives should be defined based on the availability of common outcome measures, which could be guided by our modules.
- Step 1: assemble information and select studies. Patients should be included within the same time period, or, at least, with diagnosis performed with same criteria. Age should be accounted for as a determinant of disability (e.g., age range in inclusion criteria, stratification of statistical analyses). Participating centres should contribute with populations of similar size, or, at least, specific statistical methods should be considered to reduce between-centre variability (e.g., propensity score matching).
- Step 2: define variables and evaluate harmonization potential. The study duration should be set depending on the primary outcomes (e.g., 2 years could be reasonable for relapses, but not enough for EDSS progression).
- **Step 3: process data**. Based on the previous steps, possible confounding should be evaluated, and, if necessary, should be considered statistically (e.g., age could be included as a covariate in the statistical models).

- Step 4: estimate the quality of harmonized dataset(s) generated. DerSimonian and Laird random-effects method could be used to evaluate the between-population variance in future studies.
- **Step 5: disseminate and preserve final harmonization products**. Harmonization analyses should be made available (e.g., as supplementary material).

Unfortunately, MRI measures, neuropsychological tests, and PROMs were included in a limited number of studies, and, thus, it is difficult to draw final conclusions. In the future, more detailed outcome measures could be used also for stratifying study populations, further improving the homogeneity. An additional limitation of this study is that we only analysed results from RIReMS centres, limiting results' generalisability, also considering the possible use of different tools/scales in other countries/languages. Not least, we did not evaluate heterogeneity in acquiring different outcome measures (e.g., methods for quantification of brain atrophy), which would need more indepth evaluation in future studies. However, on this ground, further cooperation between different national and international centres should be implemented (Pugliatti et al., 2012). Also, we performed retrospective analyses on already-published data to reflect the general practice of clinical registries, where data is collected locally and then analysed retrospectively (Glaser et al., 2019; Trojano et al., 2019). In a previous study, Salter and colleagues investigated retrospective harmonization of three MS registries from three different areas (Germany, North America, and the UK), with ad-hoc data collection for employment status (Salter et al., 2020); by contrast, we investigated heterogeneity at the level of MS centres (which then contribute to formation of MS registries), by evaluating common sets of demographics and clinical features, cross-sectionally and longitudinally.

In conclusion, we suggested modular sets of biomarkers for data collection in clinical practice and for research purposes, that, if acquired by multiple MS centres, could improve data harmonization and sharing. Also, by analysing the heterogeneity of previous RIReMS papers retrospectively, we developed specific indications for future study design and data analysis. Common sets of biomarkers, along with improved study design and data analysis, will increase our chance to answer, more reliably, complex research questions with real-word data, ultimately accelerating MS research.

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Figure 1. Modular sets of outcome measures

Figure shows modular sets of outcome measures. Basic modules for diagnosis and for follow-up consisted of outcome measures recorded by all participating centres at the time of diagnosis, and, then, at least annually, during follow-up. The detailed module included outcome measures recorded in most participating centres in the presence of specific clinical indications. The optional extension module included outcome measures recorded in some participating centres in the presence of specific clinical indications.

Basic module at diagnosis

Demographics: date of birth, gender Clinical features: date of onset, symptoms at onset, disease phenotype MRI features: T2, T1, and Gad-enhancing lesions in brain and cervical spinal cord (less commonly in thoracic spinal cord) Concomitant diseases and treatments Basic module at follow-up (at least annually)

Clinical features: disease phenotype Relapses: date of relapse, symptoms, recovery, steroid use Disability: FSS and EDSS MRI features: T2, and T1 lesions in brain Treatments: DMTs, symptomatic treatments, adverse events Concomitant diseases and treatments



Figure 2. Heterogeneity of demographic features.

Forest plots show heterogeneity of age at onset (**a**), and female percent (**b**) between studies including RRMS patients, estimated using the DerSimonian and Laird random-effects method. Results are presented as effect size, 95% confidence intervals (95%CI), and I².

		Age at
а.		onset (95% CI)
Bajrami, 2018	×	29.20 (28.08, 30.32)
Bisecco, 2018	+	27.60 (24.73, 30.47)
Calabrese, 2017	 =	28.80 (26.90, 30.70)
Cocozza, 2019	+	26.30 (24.37, 28.23)
Di Gregorio, 2018	-	24.80 (22.59, 27.01)
Gajofatto, 2014	-	24.80 (22.82, 26.78)
Lanzillo, 2017	+	27.19 (25.89, 28.49)
Mallucchi, 2017	*	27.00 (23.18, 30.82)
Mallucci, 2018	+	27.90 (25.76, 30.04)
Manni, 2019		25.50 (24.54, 26.46)
Megna, 2019		30.00 (29.07, 30.93)
Moccia, 2016	•	29.00 (28.13, 29.87)
Moccia, 2019		30.40 (28.96, 31.84)
Realmuto, 2018	+	24.50 (22.68, 26.32)
Scalfari, 2018		24.50 (23.76, 25.24)
Vercellino, 2009	•	29.10 (28.19, 30.01)
DL Overall (l ² = 91.7%)	۵	27.34 (26.18, 28.49)
	10 20 30 40	50

b.		Female percent (95% CI)
Bajrami, 2018		67.00 (56.63, 77.37)
Bisecco, 2018		64.40 (52.18, 76.62)
Buscarinu, 2017		68.10 (48.62, 87.58)
Calabrese, 2017		69.20 (54.71, 83.69)
Cocozza, 2019		63.50 (52.53, 74.47)
Di Filippo, 2014		67.90 (50.61, 85.19)
Di Gregorio, 2018		78.90 (68.31, 89.49)
Gajofatto, 2014	-	73.60 (64.34, 82.86)
Mallucchi, 2017		50.00 (25.50, 74.50)
Mallucci, 2018		76.00 (67.03, 84.97)
Manni, 2019	-	58.20 (52.94, 63.46)
Marastoni, 2019		60.50 (44.96, 76.04)
Megna, 2019		60.20 (53.17, 67.23)
Moccia, 2016		63.80 (56.23, 71.37)
Moccia, 2019	- <u>+</u>	70.00 (58.40, 81.60)
Realmuto, 2018		68.80 (55.26, 82.34)
Scalfari, 2018		60.00 (53.51, 66.49)
Vercellino, 2009	+	71.30 (66.21, 76.39)
DL Overall ($I^2 = 51.7\%$)	0	66.42 (63.08, 69.76)

20 40 60 80 100

Figure 3. Heterogeneity of baseline EDSS.

Forest plots show heterogeneity of baseline EDSS between studies, in subgroups of different age ranges, estimated using the DerSimonian and Laird random-effects method. Results are presented as effect size, 95% confidence intervals (95%CI), and I².

	EDSS (95% CI)
Age < 36.1 years	
Calabrese, 2017	 1.50 (1.35, 1.65)
Cocozza, 2019	★ 3.30 (3.00, 3.60)
Di Filippo, 2014	★ 1.38 (1.11, 1.65)
Di Gregorio, 2018	1.90 (1.51, 2.29)
Mallucchi, 2017	2.60 (1.87, 3.33)
Mallucci, 2018	✤ ! 2.10 (1.78, 2.42)
Megna, 2019	 2.40 (2.29, 2.51)
Moccia, 2016	 1.80 (1.74, 1.86)
Realmuto, 2018	2.06 (1.63, 2.49)
Scalfari, 2018	▲ 1.40 (1.20, 1.60)
Vercellino, 2009	 1.50 (1.37, 1.63)
DL Subtotal $(I^2 = 96.3\%)$	♦ 1.96 (1.69, 2.24)
Are 20.0.41.1 veget	
Age 36.8-41.1 years	
Bajrami, 2018	★ 2.50 (2.24, 2.76)
Bisecco, 2018	2.50 (2.07, 2.93)
Coghe, 2018	
Della Corte, 2018	✤ 2.75 (2.46, 3.04)
Gaetani, 2019	2.00 (1.56, 2.44)
Gajofatto, 2014	3.25 (2.76, 3.74)
Lanzillo, 2017	★ 3.40 (3.24, 3.56)
Manni, 2019	3.00 (2.84, 3.16)
Salemi, 2019	2.25 (1.66, 2.84)
DL Subtotal ($I^2 = 90.1\%$)	O 2.70 (2.39, 3.01)
Age > 41.4 years	
Camerota, 2017	★ 5.70 (5.44, 5.96)
Carotenuto, 2019	4.25 (3.79, 4.71)
De Biasi, 2016	2.60 (2.23, 2.97)
De Sire, 2020	★ 6.40 (6.15, 6.65)
Ferraro, 2019	 6.30 (6.19, 6.41)
Lipp. 2020	4.00 (3.91, 4.09)
Lorefice, 2019	2.80 (2.53, 3.07)
Moccia 2019	270(245,295)
Paolicelli 2016	■ 6 70 (6 49 6 91)
Pellegring 2018	5 50 (4 79 6 21)
Sola 2010	= 2.06(1.01, 2.21)
Zaccololla 2012	$ = \frac{2.00(1.31, 2.21)}{3.50(2.22, 2.77)} $
2000000000000000000000000000000000000	
DL Subiotal (I = 99.7%)	4.37 (3.40, 5.35)
DL Overall (l ² = 99.7%)	3.08 (2.50, 3.65)

Figure 4. Heterogeneity of longitudinal outcomes.

Forest plots show heterogeneity of yearly rates in relapse occurrence (**a**), and EDSS progression between longitudinal studies (**b**), in subgroups of different study durations, estimated using the DerSimonian and Laird random-effects method. Results are presented as effect size, 95% confidence intervals (95%CI), and I².



		EDSS
b.		progression (95% CI)
Study duration ≤2 years		
Bajrami, 2018	-	1.25 (-1.20, 3.70)
Calabrese, 2017		3.85 (-2.19, 9.89)
Mallucci, 2018	*	10.50 (4.06, 16.94)
DL Subtotal $(I^2 = 71.9\%)$	$\langle \rangle$	4.59 (-0.74, 9.92)
Study duration >2 years		
Cocozza, 2019		4.50 (-0.22, 9.22)
Lanzillo, 2017	. .	7.63 (3.74, 11.52)
Moccia, 2016		4.32 (1.12, 7.52)
DL Subtotal $(I^2 = 0.0\%)$	\diamond	5.41 (3.22, 7.60)
DL Overall (l ² = 59.0%)	\diamond	4.81 (2.21, 7.42)
-	10 0 10 20	30

Figure 5. Maelstrom Research guidelines with MS-specific recommendations.

Figure shows MS-specific recommendations for rigorous retrospective data harmonization, in line

with the Maelstrom Research guidelines.



Table 1. Studies included for heterogeneity analyses.

Table shows studies included for heterogeneity analyses, with details on study design, demographics and clinical features.

Reference	Disease	Sample	Study	Age at	Female	EDSS	T25FWT	9НРТ	Relapses	EDSS	Brain	GM	SDMT	PASAT3	BDI	FSS	MFIS
	subtype	Size	Duration	onset						progression	volume	volume					
		(n)	(years)	(years)	(%)		(seconds)	(seconds)	(%/year)	(%/year)	(mL)	(mL)					
Bajrami et al.	RRMS	79	2.0	29.2±5.1	67.0	2.5±1.1			14.6	1.3							
J Neurol 2018																	
Bisecco et al.	RRMS	59		27.6±11.2	64.4	2.5±1.7					1519.0±80.0	789.0±62.0					
Brain Imaging Behav 2018																	
Buscarinu et al.	RRMS	22			68.1												
Mult Scler 2017																	
Calabrese et al.	RRMS	39	2.0	28.8±6.0	69.2	1.5±0.4			6.4	3.9							
JNNP 2017																	
Camerota et al.	PMS	14			42.8	5.7±0.5	19.7±13.5								5.5±3.4		38.5±18.4
J Neurol Sci 2017																	
Carotenuto et al.	Mixed	55		30.2±13.1	69.0	4.2±1.7	13.3±10	29.9±16.4					38.9±16.0	26.8±20.7	14.4±10.3		36.9±18.7
Mult Scler Relat Disord 2019																	
Cocozza et al. Neuroradiology	RRMS	74	3.6	26.3±8.4	63.5	3.3±1.3				4.5		763.2±61.2					
2019																	
Coghe et al.	Mixed	48			68.7	2.4±1.5											
Mult Scler Relat Disord 2018																	
De Biasi et al.	Mixed	165		33.8±9.5	74.0	2.6±2.4											
Front Immunol 2016																	
De Sire et al.	Mixed	10		35.0±7.4	60.0	6.4±0.4											
NeuroRehabiltation 2019																	
Della Corte et al.	Mixed	147		27.7±10.1	72.7	2.7±1.7							35.3±14.4	34.8±13.7		32.4±16.0	
Neurol Sci 2018																	
Di Filippo et al.	RRMS	28			67.9	1.3±0.7											
Eur J Neurol 2014																	

Di Gregorio et al.	RRMS	57	2.1	24.8±8.5	78.9	1.9±1.5											
J Neurol 2018																	
Ferraro et al.	PMS	70	4.6	39.2±2.9	70.0	6.3±0.4											
Acta Neurol Scand 2019																	
Gaetani et al.	Mixed	28		36.5±10.8	60.0	2.0±1.2							42.6±12	43±14.1			
J Neurol 2019																	
Gajofatto et al.	RRMS	87	1.5	24.8±9.4	73.6	3.2±2.3			5.3								
Eur. Neurol 2014																	
Lanzillo et al.	RRMS	179	3.0	27.1±8.9		3.4±1.0			7.8	7.6							
Acta Neurol Scand 2017																	
Lipp et al.	Mixed	131		32.1±7.5	64.0	4.0±0.5	8.5±9.8	25.4±11.7			1173.5±115.9	594.4±63.2		39.9±14.0 12	2.5±10.4		39.6±20.7
Neuroimage 2020																	
Lorefice et al.	Mixed	240		29±7.4	72.0	2.8±2.1					1444.8±85.0	763.1±63.7					
Mult Scler 2019																	
Mallucchi et al.	RRMS	16	1.5	27±7.8	50.0	2.6±1.5											
Neurol Ther 2017																	
Mallucci et al.	RRMS	87	2.0	27.9±10.2	76.0	2.1±1.5				10.5							
J Neurol 2018																	
Manni et al.	RRMS	338		25.5±9	58.2	3.0±1.5											
Front Immunol 2019																	
Marastoni et al.	RRMS	38	2.0		60.5				10.0								
Front Immunol 2019																	
Megna et al. J Neurol 2019	RRMS	186		30.0±6.5	60.2	2.4±0.8							48.4±14.6	35.4±13.3			
Moccia et al.	RRMS	155	10.0	29.0±5.5	63.8	1.8±0.4			10.0	4.3			30.6±14.6	28.5±10.6			
Mult Scler 2016																	
Moccia et al.	RRMS	60	0.5	30.4±5.7	70.0	2.7±1.0			13.2					14	4.1±10.2	27.5±10.2	
Ther Adv Neurol Disord 2019																	
Paolicelli et al.	Mixed	102		29.6±9.2	49.0	6.7±1.1	28.1±17.2										
J Clin Pharmacol 2016																	
Pellegrino et al.	Mixed	11			81.0	5.5±1.2		47.8±24.2							4	47.2±11.7	
Sci Rep 2018																	
Realmuto et al.	RRMS	45		24.5±6.2	68.8	2.0±1.4							47.4±10.7		(61.0±35.1	

Arch Clin Neuropsychol 2018

Salemi et al.	Mixed	9	0.2	32.1±7.7	66.0	2.2±0.9	8.0±1.6	39±13.7	42.3±13.1
J Neural Transm 2019									
Scalfari et al.	RRMS	219	7.9	24.5±5.6	60.0	1.4±1.5			
Neurology 2018									
Sola et al.	Mixed	149		44.4±6.7	60.0	2.0±0.9			
Mult Scler 2010									
Vercellino et al.	RRMS	304	3.0	29.1±8.1	71.3	1.5±1.1	19.1		
Acta Neurol Scand 2009									
Zoccolella et al.	Mixed	217		32.1±13.5	65.4	3.5±2.0			
J Neurol 2012									